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# A generalization of almost sure local limit theorem of uniform empirical process

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## Abstract

Let  $\{X_n; n \geq 1\}$  be a sequence of independent and identically distributed  $U[0, 1]$ -distributed random variables. In this paper, we are concerned with the almost sure local central limit theorem of  $\|F_n\|$  and  $\sup_{0 \leq t \leq 1} F_n(t)$ , and some corresponding results are derived.

**MSC:** 60E15; 60F15

**Keywords:** almost sure central limit theorem; almost sure local central limit theorem; uniform empirical process

## 1 Introduction

Throughout this paper, let  $\{X_n; n \geq 1\}$  be a sequence of independent and identically distributed  $U[0, 1]$ -distributed random variables and put  $S_n = \sum_{k=1}^n X_k$ . Define the uniform empirical process  $F_n(t) = n^{-\frac{1}{2}} \sum_{i=1}^n (I_{\{X_i \leq t\}} - t)$ ,  $0 \leq t \leq 1$ ,  $\|F_n\| = \sup_{0 \leq t \leq 1} |F_n(t)|$ . It is well known that there has been recently a lively interest in probability theory concerning almost sure versions of classical limit theorems. The prototype of such a theorem is the almost sure central limit theorem (ASCLT), which has the simplest form as follows:

$$\lim_{n \rightarrow \infty} \frac{1}{\log n} \sum_{k=1}^n \frac{1}{k} I \left\{ \frac{S_k}{\sqrt{k}} \leq x \right\} = \Phi(x) \quad \text{a.s.} \quad (1.1)$$

for all  $x \in R$ ; here and in the sequel,  $I\{A\}$  is the indicator function of the event  $A$  and  $\Phi(x)$  stands for the standard normal distribution function. This result was firstly proved independently by Brosamler [1] and Schatte [2] under a stronger moment condition. Since then, this type of almost sure version, which mainly dealt with logarithmic average limit theorems, has been extended in various directions.

Especially, Fahrner and Stadtmüller [3] and Cheng *et al.* [4] extended this almost sure convergence for partial sums to the case of maxima of independent and identically distributed (i.i.d.) random variables. Under some suitable conditions, they proved the following:

$$\lim_{n \rightarrow \infty} \frac{1}{\log n} \sum_{k=1}^n \frac{1}{k} I \left\{ \frac{M_k - b_k}{a_k} \leq x \right\} = G(x) \quad \text{a.s.} \quad (1.2)$$

for all  $x \in R$ , where  $a_k > 0$  and  $b_k \in R$  satisfy

$$\lim_{k \rightarrow \infty} P\left(\frac{M_k - b_k}{a_k} \leq x\right) = G(x)$$

for any continuity point  $x$  of  $G$ .

For Gaussian sequences, Csáki and Gonchigdanzan [5] presented the validity of (1.2) for maxima of stationary Gaussian sequences under some mild conditions. Furthermore, Chen and Lin [6] extended it to non-stationary Gaussian sequences. As for some other dependent random variables, Peligrad and Shao [7] and Dudziński [8] derived some corresponding results about an almost sure central limit theorem. The almost sure central limit theorem in a joint version for log average in the case of independent and identically distributed random variables was obtained by Peng *et al.* [9], a joint version of almost sure limit theorem for log average of maxima and partial sums in the case of stationary Gaussian random variables was derived by Dudziński [10]. In this direction, an extension of almost sure central limit theory was studied by Hörmann [11].

Moreover, Wu [12–14] explored the almost sure limit theorem for product of partial sums, stable distribution and product of sums of partial sums, respectively. Zang [15] derived the almost sure limit theorem of random fields for more general weights than the usual logarithmic average. Recently, Zhang [16] established the almost sure central limit theorem for uniform empirical processes with logarithmic average. And then, under some regular conditions, a general result of almost sure central limit theorem for uniform empirical processes with general weights was derived by Zang [17] with the methodology of Hörmann [11].

On the other hand, Chung and Erdős [18] proved the following result.

**Theorem A** *Let  $X_1, X_2, \dots$  be a sequence of i.i.d. integer valued random variables with  $EX_1 = 0$ . Assume that every integer  $a$  is a possible value of  $S_k$  for all sufficiently large  $k$ . Then*

$$\lim_{n \rightarrow \infty} \frac{1}{\log T_n} \sum_{k=1}^n \frac{P(S_k = a)}{T_k} = 1 \quad a.s.,$$

where

$$T_n = \sum_{k=1}^n P(S_k = a).$$

Further, if the condition  $EX_1^2 = \sigma^2 < \infty$  is satisfied, we have  $T_k \sim \frac{\sqrt{2k}}{\sigma\sqrt{\pi}}$ , therefore

$$\lim_{n \rightarrow \infty} \frac{1}{\log n} \sum_{k=1}^n \frac{P(S_k = 0)}{\sqrt{k}} = \frac{\sqrt{2}}{\sigma\sqrt{\pi}} \quad a.s.$$

This result may be called almost sure local central limit theorem (ASLCLT), while (1.1) may be called almost sure global central limit theorem.

A more general version of this theorem was proved by Csáki *et al.* [19] with finite third moment, they derived

$$\lim_{n \rightarrow \infty} \frac{1}{\log n} \sum_{k=1}^n \frac{I\{a_k \leq S_k < b_k\}}{kP\{a_k \leq S_k < b_k\}} = 1 \quad \text{a.s.}$$

if

$$\sum_{k=1}^n \frac{\log k}{k^{\frac{3}{2}} P\{a_k \leq S_k < b_k\}} = O(\log n) \quad \text{as } n \rightarrow \infty,$$

where  $\{a_n, n \geq 1\}$  and  $\{b_n, n \geq 1\}$  are both real sequences such that  $a_n \leq 0 \leq b_n$  for  $n \geq 1$ .

In this paper, under some mild conditions, we are concerned with the almost sure local central limit theorems of  $\|F_n\|$  and  $\sup_{0 \leq t \leq 1} F_n(t)$ , which is inspired by the above result, especially Csáki *et al.* [19] and the references therein concerning the almost sure local central limit theorem.

The rest of this paper is organized as follows. In Section 2, a generalization of almost sure local central limit theorem of uniform empirical process is formulated. In Section 3, proofs of our main results are established. In Section 4, the paper is concluded and some statistical applications for future research are outlined.

## 2 Main results

In this section, let the real-valued sequences  $\{u_n, n \geq 1\}$ ,  $\{v_n, n \geq 1\}$  be such that  $u_n > v_n$ , and satisfy

$$\sum_{\substack{1 \leq k < l \leq n \\ p_k p_l \neq 0}} \frac{1}{l^2 p_k p_l} = O(\log n), \tag{2.1}$$

where  $p_k = P\{v_k < \|F_k\| \leq u_k\}$ . Set

$$\alpha_k = \begin{cases} \frac{1}{p_k} I\{v_k < \|F_k\| \leq u_k\}, & p_k \neq 0, \\ 1, & p_k = 0 \end{cases}$$

and

$$\beta_k = \begin{cases} \frac{1}{q_k} I\{v_k < \sup_{0 \leq t \leq 1} F_k(t) \leq u_k\}, & q_k \neq 0, \\ 1, & q_k = 0, \end{cases}$$

where  $q_k = P\{v_k < \sup_{0 \leq t \leq 1} F_k(t) \leq u_k\}$  satisfies the corresponding condition (2.1).

**Theorem 2.1** *Let  $\{X_n; n \geq 1\}$  be a sequence of independent and identically distributed (i.i.d.) random variables. Then*

$$\lim_{n \rightarrow \infty} \frac{1}{\log n} \sum_{k=1}^n \frac{\alpha_k}{k} = 1 \quad \text{a.s.} \tag{2.2}$$

**Theorem 2.2** Let  $\{X_n; n \geq 1\}$  be a sequence of independent and identically distributed (i.i.d.) random variables. Then

$$\lim_{n \rightarrow \infty} \frac{1}{\log n} \sum_{k=1}^n \frac{\beta_k}{k} = 1 \quad a.s. \tag{2.3}$$

**Remark 2.3** We believe that condition (2.1) can be weakened through more complicated calculating procedures, so we will study it in the future work.

**3 The proofs of the main results**

In this section, we shall give some auxiliary lemmas which will be used to prove our main result. The first lemma comes from Gonchigdanzan [20].

**Lemma 3.1** Assume that  $\xi_1, \xi_2, \dots$  are random variables such that  $E\xi_i = 1$  for  $k = 1, 2, \dots$ . Then

$$\lim_{n \rightarrow \infty} \frac{1}{\log n} E\left(\sum_{k=1}^n \frac{\xi_k}{k}\right) = 1.$$

Furthermore, if  $\xi_k \geq 0$  for  $k \geq 1$  and

$$\text{Var}\left(\sum_{k=1}^n \frac{\xi_k}{k}\right) \ll (\log n)^{2-\varepsilon}$$

for some  $\varepsilon > 0$  and large enough  $n$ , then

$$\lim_{n \rightarrow \infty} \frac{1}{\log n} \sum_{k=1}^n \frac{\alpha_k}{k} = 1 \quad a.s.$$

Here and in the sequel,  $a_n \ll b_n$  denotes  $\limsup_{n \rightarrow \infty} |a_n/b_n| < \infty$ .

**Lemma 3.2** If  $X_1, X_2, \dots$  are i.i.d. random variables with common distributed function  $F$ . Denote  $D_n = \sup_x |\frac{1}{n} \sum_{i=1}^n I\{X_i \leq x\} - F(x)|$ , then there exist positive constants  $c_0$  and  $c$  such that, for all  $n$ , all  $F$ , and all positive  $r$ , one has

$$1 - P(D_n < r/n^{1/2}) < c_0 e^{-cr^2}.$$

*Proof* This lemma is due to Kiefer and Wolfowitz [21]. □

Now, we give the proofs of our main results.

*Proof of Theorem 2.1* Denote

$$F_{k,l}(t) = l^{-\frac{1}{2}} \sum_{i=k+1}^l (I_{\{X_i \leq t\}} - t), \quad 0 \leq t \leq 1,$$

$$\|F_{k,l}\| = \sup_{0 \leq t \leq 1} |F_{k,l}(t)|.$$

Firstly, by Lemma 3.1, it only needs to verify

$$\text{Var}\left(\sum_{k=1}^n \frac{\alpha_k}{k}\right) \ll (\log n)^{2-\varepsilon}.$$

Note that

$$\text{Var}\left(\sum_{k=1}^n \frac{\alpha_k}{k}\right) = \sum_{k=1}^n \frac{\text{Var}(\alpha_k)}{k^2} + 2 \sum_{1 \leq k < l \leq n} \frac{\text{Cov}(\alpha_k, \alpha_l)}{kl}.$$

Furthermore, observe that if  $p_k = 0$ , then

$$\text{Var}(\alpha_k) = 0;$$

if  $p_k \neq 0$ , then

$$\text{Var}(\alpha_k) = \frac{1-p_k}{p_k} \leq \frac{1}{p_k},$$

and consequently

$$\sum_{k=1}^n \frac{\text{Var}(\alpha_k)}{k^2} \leq \sum_{k=1}^n \frac{1}{k^2 p_k} \leq \sum_{1 \leq k < l \leq n} \frac{1}{l^2 p_k p_l} \ll \log n;$$

if  $p_k p_l = 0$ , then

$$\text{Cov}(\alpha_k, \alpha_l) = 0;$$

if  $p_k p_l \neq 0$ , we have

$$\begin{aligned} & \text{Cov}(I\{v_k < \|F_k\| \leq u_k\}, I\{v_l < \|F_l\| \leq u_l\}) \\ & \leq |\text{Cov}(I\{\|F_k\| \leq u_k\}, I\{\|F_l\| \leq u_l\})| \\ & \quad + |\text{Cov}(I\{\|F_k\| \leq u_k\}, I\{\|F_l\| \leq v_l\})| \\ & \quad + |\text{Cov}(I\{\|F_k\| \leq v_k\}, I\{\|F_l\| \leq u_l\})| \\ & \quad + |\text{Cov}(I\{\|F_k\| \leq v_k\}, I\{\|F_l\| \leq v_l\})| \\ & =: A_1 + A_2 + A_3 + A_4. \end{aligned}$$

For  $A_1$ ,  $k \leq l$ , it follows that

$$\begin{aligned} A_1 & = |\text{Cov}(I\{\|F_k\| \leq u_k\}, I\{\|F_l\| \leq u_l\})| \\ & \leq |\text{Cov}(I\{\|F_k\| \leq u_k\}, I\{\|F_l\| \leq u_l\} - I\{\|F_{k,l}\| \leq u_l\})| \\ & \quad + |\text{Cov}(I\{\|F_k\| \leq u_k\}, I\{\|F_{k,l}\| \leq u_l\})| \\ & \ll E|I\{\|F_l\| \leq u_l\} - I\{\|F_{k,l}\| \leq u_l\}| \end{aligned}$$

$$\begin{aligned}
 &+ |\text{Cov}(I\{\|F_k\| \leq u_k\}, I\{\|F_{k,l}\| \leq u_l\})| \\
 &=: A_{11} + A_{12}.
 \end{aligned}$$

Making use of Lemma 3.2 and Fubini’s theorem, we have

$$\begin{aligned}
 A_{11} &\ll E\|\|F_l\| - \|F_{k,l}\|\| \\
 &\ll E\left| \sup_{0 \leq t \leq 1} |F_l(t)| - \sup_{0 \leq t \leq 1} |F_{k,l}(t)| \right| \\
 &\ll E \sup_{0 \leq t \leq 1} \left| |F_l(t)| - |F_{k,l}(t)| \right| \\
 &\ll E \sup_{0 \leq t \leq 1} |F_l(t) - F_{k,l}(t)| \\
 &= E \sup_{0 \leq t \leq 1} l^{-\frac{1}{2}} \left| \sum_{i=1}^k (I_{\{X_i \leq t\}} - t) \right| \\
 &= \left(\frac{k}{l}\right)^{\frac{1}{2}} E \sup_{0 \leq t \leq 1} k^{-\frac{1}{2}} \left| \sum_{i=1}^k (I_{\{X_i \leq t\}} - t) \right| \\
 &= \left(\frac{k}{l}\right)^{\frac{1}{2}} \int_0^\infty P\left\{ \sup_{0 \leq t \leq 1} k^{-\frac{1}{2}} \left| \sum_{i=1}^k (I_{\{X_i \leq t\}} - t) \right| > x \right\} dx \\
 &\ll \left(\frac{k}{l}\right)^{\frac{1}{2}} \int_0^\infty e^{-cx^2} dx \\
 &\ll \left(\frac{k}{l}\right)^{\frac{1}{2}}.
 \end{aligned}$$

In virtue of the independence of  $\{X_n; n \geq 1\}$ , we have  $A_{12} = 0$ . Then

$$A_1 \ll \left(\frac{k}{l}\right)^{\frac{1}{2}}.$$

Furthermore, by similar reasoning, we have

$$A_i \ll \frac{k}{l}, \quad i = 2, 3, 4,$$

and therefore,

$$A_2 + A_3 + A_4 \ll \frac{k}{l}.$$

Thus, according to our assumptions, we can derive

$$\begin{aligned}
 &\sum_{1 \leq k < l \leq n} \frac{\text{Cov}(I\{v_k < R_k \leq u_k\}, I\{v_l < R_l \leq u_l\})}{klp_k p_l} \\
 &\leq \sum_{1 \leq k < l \leq n} \frac{A_1 + A_2 + A_3 + A_4}{klp_k p_l}
 \end{aligned}$$

$$\begin{aligned}
&\ll \sum_{1 \leq k < l \leq n} \frac{1}{klp_k p_l} \left(\frac{k}{l}\right)^{\frac{1}{2}} \\
&= \sum_{1 \leq k < l \leq n} \frac{1}{\sqrt{k} l^{3/2} p_k p_l} \\
&= O(\log n).
\end{aligned}$$

Hence,  $\text{Var}(\sum_{k=1}^n \frac{a_k}{k}) \ll (\log n)$ . Further, we complete the proof of Theorem 2.1.  $\square$

*Proof of Theorem 2.2* The proof is similar to the above procedures, so we omit it here.  $\square$

#### 4 Concluding remarks

In this paper, we are concerned with the limit theory of uniform empirical process. A generalization of almost sure local central limit theorem of uniform empirical process has been established.

Some statistical applications related to our main result deserve further investigation. By virtue of being a new approach of testing based on ASCLT, Thangavelu [22] investigated hypothesis testing via the estimation of quantiles of the distribution of the concerned statistics. Based on the theorem on ASCLT for rank statistics, he also proposed a small-sample approximation for the two-sample nonparametric Behrens-Fisher problem. These statistical applications concerning our work will be discussed in the future work.

#### Competing interests

The author declares that they have no competing interests.

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