# **RIO-TYPE INEQUALITY FOR THE EXPECTATION OF PRODUCTS OF RANDOM VARIABLES**

### B. L. S. PRAKASA RAO

Received 26 July 2004

We develop an inequality for the expectation of a product of n random variables generalizing the recent work of Dedecker and Doukhan (2003) and the earlier results of Rio (1993).

# 1. Introduction

Let  $(\Omega, \mathcal{F}, P)$  be a probability space and let (X, Y) be a bivariate random vector defined on it. Suppose that  $E(X^2) < \infty$  and  $E(Y^2) < \infty$ . Hoeffding proved that

$$\operatorname{Cov}(X,Y) = \int_{\mathbb{R}^2} \left[ P(X \le x, \ Y \le y) - P(X \le x) P(Y \le y) \right] dx \, dy.$$
(1.1)

In [5], Lehmann gave a simple proof of this identity and used it in his study of some concepts of dependence. This identity was generalized to functions h(X) and g(Y) with  $E[h^2(X)] < \infty$  and  $E[g^2(Y)] < \infty$  and with finite derivatives  $h'(\cdot)$  and  $g'(\cdot)$  by Newman [6]. Multidimensional versions of these results were proved by Block and Fang [1], Yu [13], and more recently by Prakasa Rao [7]. Related covariance identities for exponential and other distributions are given by Prakasa Rao in [9, 10].

Suppose that  $\mathcal{M}$  is a sub- $\sigma$ -algebra of  $\mathcal{F}$  and Y is measurable with respect to  $\mathcal{M}$ . Let  $\sigma(X)$  be the sub- $\sigma$ -algebra generated by the random variable X. Define

$$\alpha(\mathcal{M}, X) = \sup\left\{ \left| P(A \cap B) - P(A)P(B) \right|, A \in \mathcal{M}, B \in \sigma(X) \right\}.$$
(1.2)

Define

$$Q_X(u) = \inf \{ x : P(|X| > x) \le u \},\$$

$$G_X(s) = \inf \{ z : \int_0^z Q_X(t) dt \ge s \},\$$

$$H_{X,Y}(s) = \inf \{ t : E(|X|I_{[|Y|>t]}) \le s \}.$$
(1.3)

Copyright © 2005 Hindawi Publishing Corporation Journal of Inequalities and Applications 2005:1 (2005) 7–14 DOI: 10.1155/JIA.2005.7

# 8 Rio-type inequality

Rio [11] proved that

$$\left|\operatorname{Cov}(X,Y)\right| \le 2 \int_0^{\alpha(\mathcal{M},X)/2} Q_Y(u) Q_X(u) du.$$
(1.4)

Related results are given in [12, page 9]. These results were generalized by Bradley [2] for a strong-mixing process and by Prakasa Rao [8] for *r*th-order joint cumulant under *r*th-order strong mixing. In a recent work, Dedecker and Doukhan [3] proved that

$$\left| E(XY) \right| \leq \int_{0}^{\|E(X|\mathcal{M})\|_{1}} H_{X,Y}(t) dt \leq \int_{0}^{\|E(X|\mathcal{M})\|_{1}} Q_{Y} oG_{X}(t) dt$$
(1.5)

and obtained an improved version of the above inequality. If  $X_i$ ,  $1 \le i \le n$ , are positive-valued random variables, it is easy to see that

$$E(X_1X_2\cdots X_n) \leq \int_0^1 Q_{X_1}(u)Q_{X_2}(u)\cdots Q_{X_n}(u)du.$$
(1.6)

For a proof, see [12, Lemma 2.1, page 35].

We now obtain an improved version of the above inequality following the techniques of Dedecker and Doukhan [3] and Block and Fang [1].

# 2. Main result

Let  $\{X_i, 1 \le i \le n\}$  be a sequence of nonnegative random variables defined on a probability space  $(\Omega, \mathcal{F}, P)$ . Then the random variable  $X_i$  can be represented in the form

$$X_{i} = \int_{0}^{\infty} I_{(x_{i},\infty)}(X_{i}) dx_{i}, \qquad (2.1)$$

where

$$I_{(x_i,\infty)}(X_i) = \begin{cases} 1 & \text{if } X_i > x_i, \\ 0 & \text{if } X_i \le x_i. \end{cases}$$
(2.2)

Hence

$$E(X_{1}X_{2}\cdots X_{n}) = E\left[X_{1}\Pi_{i=2}^{n}\int_{0}^{\infty}I_{(x_{i},\infty)}(X_{i})dx_{i}\right]$$
  
$$=\int_{\mathbb{R}^{n-1}_{+}}E[X_{1}\Pi_{i=2}^{n}I_{(x_{i},\infty)}(X_{i})]dx_{2}\cdots dx_{n}$$
  
$$=\int_{\mathbb{R}^{n-1}_{+}}E[X_{1}I_{[X_{i}>x_{i},2\leq i\leq n]}(X_{2},\ldots,X_{n})]dx_{2}\cdots dx_{n}$$
  
(2.3)

by the Fubini's theorem, where  $\mathbb{R}^{n-1}_+ = \{(x_2, \dots, x_n) : x_i \ge 0, 2 \le i \le n\}$ . Observe that

$$E(X_1 I_{[X_i > x_i, 2 \le i \le n]}(X_2, \dots, X_n)) \le \min(E[X_1], E(X_1 I_{[X_i > x_i, 2 \le i \le n]}(X_2, \dots, X_n)))$$
(2.4)

and hence

$$E(X_1X_2\cdots X_n) \leq \int_{\mathbb{R}^{n-1}_+} \left\{ \int_0^{EX_1} \chi_{(E[X_1I_{[X_i>x_i,2\leq i\leq n]}(X_2,\dots,X_n)]>u)}(u) du \right\} dx_2\cdots dx_n.$$
(2.5)

Here  $\chi_A(\cdot)$  denotes the indicator function of the set *A*. Let

$$g_{X_1}(x_2,\ldots,x_n) = E[X_1 I_{[X_i > x_i, 2 \le i \le n]}(X_2,\ldots,X_n)].$$
(2.6)

Then

$$E(X_{1}X_{2}\cdots X_{n}) \leq \int_{\mathbb{R}^{n-1}_{+}} \left\{ \int_{0}^{EX_{1}} \chi_{[g_{X_{1}}(x_{2},...,x_{n})>u]}(u) du \right\} dx_{2}\cdots dx_{n}$$

$$= \int_{0}^{E(X_{1})} \left\{ \int_{[(x_{2},...,x_{n}):g_{X_{1}}(x_{2},...,x_{n})>u]} 1 dx_{2}\cdots dx_{n} \right\} du.$$
(2.7)

Let

$$H_{X_1,X_2,...,X_n}(u) = \lambda [(x_2,...,x_n) : g_{X_1}(x_2,...,x_n) > u],$$
(2.8)

where  $\lambda$  is the Lebesgue measure on the space  $\mathbb{R}^{n-1}_+$ . Hence

$$E(X_1X_2\cdots X_n) \leq \int_0^{E(X_1)} H_{X_1,X_2,\dots,X_n}(u) du.$$
 (2.9)

Observe that

$$g_{X_1}(x_2,\ldots,x_n) = E[X_1I_{[X_i > x_i, 2 \le i \le n]}(X_2,\ldots,X_n)] \le \int_0^{E[I_{[X_i > x_i, 2 \le i \le n]}(X_2,\ldots,X_n)]} Q_{X_1}(u) du$$
(2.10)

from the Fréchet's inequality [4]. Here  $Q_{X_1}(\cdot)$  is the generalized inverse of the function  $T_{X_1}(x) = P(X_1 > x)$  as defined earlier. Let

$$M_{X_1}(y) = \int_0^y Q_{X_1}(t) dt.$$
 (2.11)

Observe that  $M_{X_1}(\cdot)$  is nondecreasing in y. Let  $G_{X_1}(u) = \inf\{z : M_{X_1}(z) \ge u\}$  as defined earlier. Let

$$T_{X_2,\dots,X_n}(x_2,\dots,x_n) = P(X_i > x_i, \ 2 \le i \le n).$$
(2.12)

Note that

$$g_{X_1}(x_2,...,x_n) \le M_{X_1}(E(I_{[X_i > x_i, 2 \le i \le n]}(X_2,...,X_n))),$$
  

$$g_{X_1}(x_2,...,x_n) > u \Longrightarrow M_{X_1}(E(I_{[X_i > x_i, 2 \le i \le n]}(X_2,...,X_n))) > u$$
  

$$\implies E(I_{[X_i > x_i, 2 \le i \le n]}(X_2,...,X_n)) > G_{X_1}(u)$$
  

$$\implies P[X_i > x_i, 2 \le i \le n] > G_{X_1}(u).$$
(2.13)

### 10 Rio-type inequality

Hence the set

$$[(x_2,...,x_n) \in \mathbb{R}^{n-1}_+ : g_{X_1}(x_2,...,x_n) > u]$$
(2.14)

is contained in the set

$$[(x_2,...,x_n) \in \mathbb{R}^{n-1}_+ : P(X_i > x_i, 2 \le i \le n) > G_{X_1}(u)].$$
(2.15)

In particular, it follows that the Lebesgue measure of the former set is less than or equal to that of the latter. Let

$$Q_{X_2,\dots,X_n}^*(G_{X_1}(u)) \tag{2.16}$$

denote the Lebesgue measure of the set (2.15).

Then

$$H_{X_1, X_2, \dots, X_n}(u) \le Q^*_{X_2, \dots, X_n}(G_{X_1}(u))$$
(2.17)

for all  $0 \le u \le 1$ . Hence

$$E(X_1X_2\cdots X_n) \leq \int_0^{E(X_1)} Q^*_{X_2,\dots,X_n}(G_{X_1}(u)) du.$$
 (2.18)

We have proved the following inequality.

THEOREM 2.1. Let  $X_i$ ,  $1 \le i \le n$ , be nonnegative random variables defined on a probability space  $(\Omega, \mathcal{F}, P)$ . Then

$$E(X_1X_2\cdots X_n) \leq \int_0^{E(X_1)} H_{X_1,X_2,\dots,X_n}(u) du \leq \int_0^{E(X_1)} Q^*_{X_2,\dots,X_n} oG_{X_1}(u) du,$$
(2.19)

where the functions  $H, Q^*$ , and G are as defined earlier.

# 3. Applications

We now suppose that the random variables  $\{X_i, 1 \le i \le n\}$  are arbitrary but with

$$E\left|X_{1}X_{2}\cdots X_{n}\right| < \infty. \tag{3.1}$$

Define

$$g_{X_1}(x_2,...,x_n) = E(|X_1| I_{[|X_i| > x_i, 2 \le i \le n]}(X_2,...,X_n)),$$
  

$$H_{X_1,X_2,...,X_n}(u) = \lambda[(x_2,...,x_n) : g_{X_1}(x_2,...,x_n) \le u],$$
  

$$T_{X_2,...,X_n}(x_2,...,x_n) = P(|X_i| > x_i, 2 \le i \le n),$$
(3.2)

and define  $M_{X_1}(\cdot), Q_{X_1}(\cdot), Q^*_{X_2,\dots,X_n}$ , and  $G_{X_1}$  accordingly. The following theorem follows by arguments analogous to those given in Section 2.

THEOREM 3.1. Let  $X_i$ ,  $1 \le i \le n$ , be arbitrary random variables defined on a probability space  $(\Omega, \mathcal{F}, P)$ . Then

$$E(|X_1X_2\cdots X_n|) \le \int_0^{E(|X_1|)} H_{X_1,X_2,\dots,X_n}(u) du \le \int_0^{E(|X_1|)} Q^*_{X_2,\dots,X_n} oG_{X_1}(u) du, \qquad (3.3)$$

where the functions  $H, Q^*$ , and G are as defined above.

In particular, for n = 2, we have

$$E(|X_1X_2|) \le \int_0^{E(|X_1|)} H_{X_1,X_2}(u) du \le \int_0^{E(|X_1|)} Q_{X_2} oG_{X_1}(u) du$$
(3.4)

since  $Q_X^* = Q_X$  for any univariate random variable *X*. Furthermore,

$$G_{X_1-E(X_1)}(u) \ge G_{X_1}\left(\frac{u}{2}\right), \quad 0 \le u \le 1$$
 (3.5)

(cf. [3]). Hence

$$E[|X_1X_2|] \le \int_0^{G_{X_1}^{-1}(E(|X_1|)/2)} Q_{X_2}(u)Q_{X_1}(u)du.$$
(3.6)

Therefore, for any two functions  $f_i(\cdot)$ , i = 1, 2, with  $f_i(0) = 0$  such that  $E|f_1(X_1)f_2(X_2)| < \infty$ , we obtain that

$$E[|f_1(X_1)f_2(X_2)|] \le \int_0^{G_{f_1(X_1)}^{-1}(E(|f_1(X_1)|)/2)} Q_{f_2(X_2)}(u)Q_{f_1(X_1)}(u)du.$$
(3.7)

Applying Theorem 3.1 for the random variables  $X_1 - E(X_1), X_2, \dots, X_n$ , we get that

$$E[|(X_1 - E(X_1))X_2 \cdots X_n|] \le \int_0^{E(|X_1 - E(X_1)|)} Q^*_{X_2, \dots, X_n} oG_{X_1 - E(X_1)}(u) du.$$
(3.8)

But

$$G_{X_1-E(X_1)}(u) \ge G_{X_1}\left(\frac{u}{2}\right), \quad u \ge 0$$
 (3.9)

(cf. [3]). Hence

$$E[|(X_1 - E(X_1))X_2 \cdots X_n|] \le \int_0^{E(|X_1 - E(X_1)|)/2} Q^*_{X_2, \dots, X_n} oG_{X_1}(u) du.$$
(3.10)

Observing that  $G_{X_1}(\cdot)$  is the inverse of the function  $M_{X_1}(y) = \int_0^y Q_{X_1}(t) dt$ , it follows that

$$E[|(X_1 - E(X_1))X_2 \cdots X_n|] \le \int_0^{G_{X_1}^{-1}(E(|X_1 - E(X_1)|)/2)} Q_{X_2,\dots,X_n}^*(u)Q_{X_1}(u)du.$$
(3.11)

Hence we have the following result.

#### 12 Rio-type inequality

THEOREM 3.2. Let  $X_i$ ,  $1 \le i \le n$ , be arbitrary random variables defined on a probability space  $(\Omega, \mathcal{F}, P)$  with  $E|X_1| < \infty$  and  $E|X_1X_2 \cdots X_n| < \infty$ . Then (3.11) holds.

Observe that  $Q_X^* = Q_X$  for any univariate random variable *X*. Let n = 2 in Theorem 3.2. Then  $Q_{X_2}^* = Q_{X_2}$  and the above result reduces to

$$E[|(X_1 - E(X_1))X_2|] \le \int_0^{G_{X_1}^{-1}(E(|X_1 - E(X_1)|)/2)} Q_{X_2}(u)Q_{X_1}(u)du.$$
(3.12)

As a further consequence, we get that

$$E[|(X_1 - E(X_1))(X_2 - E(X_2))|] \le \int_0^{G_{X_1}^{-1}(E(|X_1 - E(X_1)|)/2)} Q_{X_2 - E(X_2)}(u)Q_{X_1}(u)du. \quad (3.13)$$

Since

$$Q_{X_2 - E(X_2)} \le Q_{X_2} + E \left| X_2 \right|, \tag{3.14}$$

we obtain that

$$E[|(X_{1} - E(X_{1}))(X_{2} - E(X_{2}))|]$$

$$\leq \int_{0}^{G_{X_{1}}^{-1}(E(|X_{1} - E(X_{1})|)/2)} Q_{X_{2}}(u)Q_{X_{1}}(u)du + E|X_{2}|\int_{0}^{G_{X_{1}}^{-1}(E(|X_{1} - E(X_{1})|)/2)} Q_{X_{1}}(u)du.$$
(3.15)

Let

$$\alpha(X_1, X_2) = \max\left\{G_{X_1}^{-1}\left(\frac{E(|X_1 - E(X_1)|)}{2}\right), G_{X_2}^{-1}\left(\frac{E(|X_2 - E(X_2)|)}{2}\right)\right\}.$$
 (3.16)

Then it follows that

$$E[|(X_{1} - E(X_{1}))(X_{2} - E(X_{2}))|]$$

$$\leq \int_{0}^{\alpha(X_{1},X_{2})} Q_{X_{1}}(u)Q_{X_{2}}(u)du + \frac{1}{2}\left(E|X_{1}|\int_{0}^{\alpha(X_{1},X_{2})} Q_{X_{1}}(u)du + E|X_{2}|\int_{0}^{\alpha(X_{1},X_{2})} Q_{X_{2}}(u)du\right).$$
(3.17)

This inequality is different from the inequality in [12, page 9].

Let  $f_1$  and  $f_2$  be differentiable functions on  $\mathbb{R}_+$  with  $f_i(0) = 0$ . Let  $X_i$ , i = 1, 2, be non-negative random variables. Suppose that  $E[f_i^2(X_i)] < \infty$ , i = 1, 2. It is easy to see that

$$f_i(X_i) = \int_0^\infty f'_i(X_i) I_{(x_i,\infty)}(X_i) dx_i.$$
(3.18)

Then

$$E(f_{1}(X_{1})f_{2}(X_{2})) = E\left[f_{1}(X_{1})\int_{0}^{\infty} f_{2}'(X_{2})I_{(x_{2},\infty)}(X_{2})dx_{2}\right]$$
  
$$= \int_{\mathbb{R}_{+}} E[f_{1}(X_{1})f_{2}'(X_{2})I_{(x_{2},\infty)}(X_{2})]dx_{2}$$
(3.19)

B. L. S. Prakasa Rao 13

by the Fubini's theorem. Observe that

$$E(|f_{1}(X_{1})f_{2}'(X_{2})|I_{[X_{2}>x_{2}]}(X_{2})) \leq \min(E[|f_{1}(X_{1})f_{2}'(X_{2})|], E(|f_{1}(X_{1})f_{2}'(X_{2})|I_{[X_{2}>x_{2}]}(X_{2})))$$
(3.20)

and hence

$$E(f_{1}(X_{1})f_{2}(X_{2})) |$$

$$\leq \int_{\mathbb{R}^{+}} \left\{ \int_{0}^{E[|f_{1}(X_{1})f_{2}'(X_{2})|]} \chi_{(E[|f_{1}(X_{1})f_{2}'(X_{2})|I_{[X_{2}>x_{2}]}(X_{2})]>u)}(u) du \right\} dx_{2}.$$
(3.21)

Here  $\chi_A(\cdot)$  denotes the indicator function of the set *A*. Let

$$g_{f_1(X_1),f'_2(X_2)}(x_2) = E[|f_1(X_1)f'_2(X_2)|I_{[X_2>x_2]},(X_2)].$$
(3.22)

Then

$$|E(f_{1}(X_{1})f_{2}(X_{2}))| \leq \int_{\mathbb{R}_{+}} \left\{ \int_{0}^{E[|f_{1}(X_{1})f_{2}'(X_{2})|]} \chi([g_{f_{1}(X_{1}),f_{2}'(X_{2})}(x_{2})] > u)(u) du \right\} dx_{2}$$
  
$$\leq \int_{0}^{E[|f_{1}(X_{1})f_{2}'(X_{2})|]} \left\{ \int_{[x_{2}:g_{f_{1}(X_{1}),f_{2}'(X_{2})}(x_{2}) > u]} 1 dx_{2} \right\} du.$$
(3.23)

Let

$$H_{f_1(X_1), f'_2(X_2)}(u) = \inf \{ x_2 : g_{f_1(X_1), f'_2(X_2)}(x_2) \le u \}.$$
(3.24)

Then it follows that

$$\left| E(f_1(X_1)f_2(X_2)) \right| \le \int_0^{E[|f_1(X_1)f_2'(X_2)|]} H_{f_1(X_1),f_2'(X_2)}(u) du.$$
(3.25)

An analogous inequality holds by interchanging  $f_1(X_1)$  and  $f_2(X_2)$ :

$$\left| E(f_1(X_1)f_2(X_2)) \right| \leq \int_0^{E[|f_1'(X_1)f_2(X_2)|]} H_{f_1'(X_1),f_2(X_2)}(u) du.$$
(3.26)

# References

- H. W. Block and Z. B. Fang, A multivariate extension of Hoeffding's lemma, Ann. Probab. 16 (1988), no. 4, 1803–1820.
- [2] R. C. Bradley, A covariance inequality under a two-part dependence assumption, Statist. Probab. Lett. 30 (1996), no. 4, 287–293.
- [3] J. Dedecker and P. Doukhan, A new covariance inequality and applications, Stochastic Process. Appl. 106 (2003), no. 1, 63–80.
- [4] M. Fréchet, Sur la distance de deux lois de probabilité, C. R. Acad. Sci. Paris 244 (1957), 689–692 (French).
- [5] E. L. Lehmann, Some concepts of dependence, Ann. Math. Statist. 37 (1966), 1137–1153.
- [6] C. M. Newman, Normal fluctuations and the FKG inequalities, Comm. Math. Phys. 74 (1980), no. 2, 119–128.

- 14 Rio-type inequality
- B. L. S. Prakasa Rao, *Hoeffding identity, multivariance and multicorrelation*, Statistics 32 (1998), no. 1, 13–29.
- [8] \_\_\_\_\_, Bounds for rth order joint cumulant under rth order strong mixing, Statist. Probab. Lett.
   43 (1999), no. 4, 427–431.
- [9] \_\_\_\_\_, Covariance identities for exponential and related distributions, Statist. Probab. Lett. 42 (1999), no. 3, 305–311.
- [10] \_\_\_\_\_, Some covariance identities, inequalities and their applications: a review, Proc. Indian Nat. Sci. Acad. Part A 66 (2000), no. 5, 537–543.
- [11] E. Rio, Covariance inequalities for strongly mixing processes, Ann. Inst. H. Poincaré Probab. Statist. 29 (1993), no. 4, 587–597.
- [12] \_\_\_\_\_, Théorie Asymptotique des Processus Aléatoires Faiblement Dépendants, Springer, Paris, 2000.
- [13] H. Yu, A Glivenko-Cantelli lemma and weak convergence for empirical processes of associated sequences, Probab. Theory Related Fields 95 (1993), no. 3, 357–370.

B. L. S. Prakasa Rao: Theoretical Statistics and Mathematics Unit, Indian Statistical Institute, New Delhi 110 016, India

E-mail address: blsp@isid.ac.in