Research Article

Maximal Inequalities for Dependent Random Variables and Applications

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For a sequence $\{X_n, n \geq 1\}$ of dependent square integrable random variables and a sequence $\{b_n, n \geq 1\}$ of positive numbers, we establish a maximal inequality for weighted sums of dependent random variables. Applying this inequality, we obtain the almost sure convergence of $\sum_{i=1}^{n} X_i/b_i$ and $\sum_{i=1}^{n} X_i/b_n$.

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1. Introduction

Throughout this paper let $\{X_n, n \geq 1\}$ be a sequence of random variables defined on a probability space (Ω, \mathcal{F}, P) and let $\{b_n, n \geq 1\}$ be a sequence of positive numbers. We assume that there exists a sequence $\{\rho_n, n \geq 1\}$ of nonnegative constants such that

$$\sup_{k\geq 1} E(X_k X_{k+n}) \leq \rho_n, \quad \text{for } n \geq 1.$$
 (1.1)

In this paper, we establish a maximal inequality for weighted sums of the dependent random variables satisfying (1.1). Applying this inequality, we obtain under some suitable conditions on the sequence $\{\rho_n\}$ that

$$\sum_{i=1}^{n} \frac{X_i}{b_i} \text{ converges a.s. as } n \longrightarrow \infty$$
 (1.2)

and the strong law of large numbers (SLLN)

$$\frac{\sum_{i=1}^{n} X_i}{h_n} \longrightarrow 0 \text{ a.s.}$$
 (1.3)

Note that if $0 < b_n \uparrow \infty$, then (1.2) implies (1.3) by the Kronecker lemma.

For a sequence of dependent random variables satisfying (1.1), the SLLNs were established by Hu et al. [1, 2] and Lyons [3]. Lyons [3] obtained an SLLN under the conditions that $Var(X_n) = O(1)$ and $b_n = n$. Without condition $Var(X_n) = O(1)$, Hu et al. [1] obtained an SLLN, where $b_n = n$. Hu et al. [2] also obtained an SLLN for more general sequence $\{b_n\}$ $\{b_n = n \text{ is replaced by } n = O(b_n)\}$.

For other results on the SLLN for a sequence of correlated random variables, see Chandra [4], Móricz [5, 6], and Serfling [7, 8].

In this paper, we give a sufficient condition under which (1.2) and (1.3) hold. Our results (partially) improve those of Hu et al. [1, 2]. The technique used in our proof is the well-known method of subsequences. Note that the maximal inequality is used in the method of subsequences. Our maximal inequality for weighted sums of the dependent random variables satisfying (1.1) is sharper than that of Hu et al. [2].

Throughout this paper, $\log x$ denotes the natural logarithm.

2. Maximal inequalities for dependent random variables

To prove the maximal inequality for weighted sums of dependent random variables satisfying (1.1), the following lemma is needed.

Lemma 2.1. Let $\{X_n, n \ge 1\}$ be a sequence of square integrable random variables satisfying (1.1). Let $\{b_n, n \ge 1\}$ be a sequence of positive numbers such that

$$n \le Db_n \quad \forall n \ge 1 \text{ and some constant } D > 0.$$
 (2.1)

Then for all $n \ge 1$, m > n, and $\delta > 0$,

$$\sum_{i=n}^{m-1} \sum_{j=i+1}^{m} \frac{(EX_i X_j)^+}{b_i b_j} \le \frac{D^2 C_{\delta}}{\max\{(\log 2)^{\delta}, (\log n)^{\delta}\}} \sum_{k=1}^{m-n} \frac{\rho_k}{k} (1 + \log k)^{1+\delta}, \tag{2.2}$$

where $C_{\delta} = 2^{\delta+1} \max\{1, \delta^{\delta} e^{-\delta}\}.$

Proof. For simplicity of notation, let $I_{n,m} = \sum_{i=n}^{m-1} \sum_{j=i+1}^m (EX_iX_j)^+/(b_ib_j)$. Then we get by (1.1) and (2.1) that for $1 \le n < m$,

$$I_{n,m} \leq \sum_{i=n}^{m-1} \sum_{j=i+1}^{m} \frac{\rho_{j-i}}{b_i b_j}$$

$$\leq D^2 \sum_{i=n}^{m-1} \sum_{j=i+1}^{m} \frac{\rho_{j-i}}{ij}$$

$$= D^2 \sum_{k=1}^{m-n} \sum_{i=n}^{m-k} \frac{\rho_k}{i(i+k)}$$

$$= D^2 \sum_{k=1}^{m-n} \frac{\rho_k}{k} \sum_{i=n}^{m-k} \left(\frac{1}{i} - \frac{1}{i+k}\right)$$

$$\leq D^2 \sum_{k=1}^{m-n} \frac{\rho_k}{k} \sum_{i=n}^{n+k-1} \frac{1}{i}.$$
(2.3)

We next estimate $\sum_{i=n}^{n+k-1} 1/i$. If n = 1, then

$$\sum_{i=n}^{n+k-1} \frac{1}{i} = \sum_{i=1}^{k} \frac{1}{i} \le 1 + \int_{1}^{k} \frac{1}{x} dx \le 1 + \log k \le (1 + \log k)^{1+\delta}. \tag{2.4}$$

If $n \ge 2$, then

$$\sum_{i=n}^{n+k-1} \frac{1}{i} \le \int_{n-1}^{n+k-1} \frac{1}{x} dx = \log\left(1 + \frac{k}{n-1}\right) \le 2\log\left(1 + \frac{k}{n}\right). \tag{2.5}$$

The $\log(1 + k/n)$ is estimated as follows:

$$\log\left(1+\frac{k}{n}\right) \leq \begin{cases} \log\left(1+\frac{1}{\sqrt{n}}\right) \leq \frac{(\log n)^{\delta}}{(\log n)^{\delta}\sqrt{n}} \leq \frac{(2\delta)^{\delta}e^{-\delta}}{(\log n)^{\delta}} \leq (2\delta)^{\delta}e^{-\delta}\frac{(1+\log k)^{1+\delta}}{(\log n)^{\delta}}, & \text{if } 1\leq k\leq\sqrt{n}, \\ \log\left(1+\frac{k}{n}\right)\frac{(2\log k)^{\delta}}{(\log n)^{\delta}} \leq 2^{\delta}\frac{(\log k)^{1+\delta}}{(\log n)^{\delta}} \leq 2^{\delta}\frac{(1+\log k)^{1+\delta}}{(\log n)^{\delta}}, & \text{if } k>\sqrt{n}. \end{cases}$$

$$(2.6)$$

Thus, we have the desired estimate for $I_{n,m}$:

$$I_{n,m} \leq \begin{cases} D^{2} \sum_{k=1}^{m-n} \frac{\rho_{k}}{k} (1 + \log k)^{1+\delta}, & \text{if } n = 1, \\ D^{2} \sum_{k=1}^{m-n} \frac{\rho_{k}}{k} \frac{2 \max \{2^{\delta}, (2\delta)^{\delta} e^{-\delta}\}}{(\log n)^{\delta}} (1 + \log k)^{1+\delta}, & \text{if } n \geq 2, \end{cases}$$

$$(2.7)$$

$$\leq \frac{D^2 2^{\delta+1} \max\{1, \delta^{\delta} e^{-\delta}\}}{\max\{(\log 2)^{\delta}, (\log n)^{\delta}\}} \sum_{k=1}^{m-n} \frac{\rho_k}{k} (1 + \log k)^{1+\delta}.$$

The following lemma is a maximal inequality for general dependent random variables.

Lemma 2.2. Let $\{X_n, n \ge 1\}$ be a sequence of square integrable random variables. Then for all $a \ge 0$ and $n \ge 1$,

$$E\left(\max_{1\leq k\leq n}\left|\sum_{i=a+1}^{a+k}X_i\right|^2\right) \leq \left(\frac{\log 2n}{\log 2}\right)^2 \left\{\sum_{i=a+1}^{a+n}EX_i^2 + 2\sum_{i=a+1}^{a+n-1}\sum_{j=i+1}^{a+n}(EX_iX_j)^+\right\}. \tag{2.8}$$

Proof. Let $F_{a,n}$ be the joint distribution function of X_{a+1}, \ldots, X_{a+n} . Define a function g on $\{F_{a,n} : a \ge 0, n \ge 1\}$ by

$$g(F_{a,n}) = \sum_{i=a+1}^{a+n} EX_i^2 + 2\sum_{i=a+1}^{a+n-1} \sum_{j=i+1}^{a+n} (EX_i X_j)^+.$$
 (2.9)

Then we can easily obtain that for $a \ge 0$, $k \ge 1$, and $m \ge 1$,

$$g(F_{a,k}) + g(F_{a+k,m}) \le g(F_{a,k+m}).$$
 (2.10)

Moreover, we have that for all $a \ge 0$ and $n \ge 1$,

$$E\left(\sum_{i=a+1}^{a+n} X_i\right)^2 \le g(F_{a,n}). \tag{2.11}$$

By Serfling's [9] generalization of the Rademacher-Menchoff maximal inequality for orthogonal random variables,

$$E\left(\max_{1\leq k\leq n}\left|\sum_{i=a+1}^{a+k} X_i\right|^2\right) \leq \left(\frac{\log 2n}{\log 2}\right)^2 g(F_{a,n}). \tag{2.12}$$

Thus, the result is proved.

Combining Lemmas 2.1 and 2.2 gives the following maximal inequality for weighted sums of dependent random variables satisfying (1.1).

Lemma 2.3. Let $\{X_n, n \ge 1\}$ be a sequence of square integrable random variables satisfying (1.1). Let $\{b_n, n \ge 1\}$ be a sequence of positive numbers satisfying (2.1). Then for all $n \ge 1$, m > n, and $\delta > 0$,

$$E\left(\max_{n\leq i\leq m} \left| \sum_{j=n}^{i} \frac{X_{j}}{b_{j}} \right|^{2}\right)$$

$$\leq \left(\frac{\log(2(m-n+1))}{\log 2}\right)^{2} \left\{ \sum_{i=n}^{m} \frac{EX_{i}^{2}}{b_{i}^{2}} + \frac{2D^{2}C_{\delta}}{\max\{(\log 2)^{\delta}, (\log n)^{\delta}\}} \sum_{k=1}^{m-n} \frac{\rho_{k}}{k} (1 + \log k)^{1+\delta} \right\}, \tag{2.13}$$

where $C_{\delta} = 2^{\delta+1} \max\{1, \delta^{\delta} e^{-\delta}\}.$

3. Almost surely convergent series and strong laws of large numbers

In this section, we will assume that $\{X_n, n \ge 1\}$ is a sequence of square integrable random variables satisfying (1.1). A sufficient condition will be given under which (1.2) and (1.3) hold.

We first state and prove one of our main results. The proof is based on the well known method of subsequences. Our proof is similar to that of Hu et al. [2]. However, the maximal inequality (Lemma 2.3) used in the proof is sharper than that of Hu et al. [2].

Theorem 3.1. Let $\{X_n, n \ge 1\}$ be a sequence of square integrable random variables satisfying (1.1). Let $\{b_n, n \ge 1\}$ be a sequence of positive numbers satisfying (2.1). Suppose that the following conditions hold:

(i)
$$\sum_{n=1}^{\infty} (\log n)^2 E X_n^2 / b_n^2 < \infty$$
,

(ii)
$$\sum_{n=1}^{\infty} (\log n)^{4+\delta} \rho_n / n < \infty$$
 for some $\delta > 0$.

Then (1.2) holds. Furthermore, if $0 < b_n \uparrow \infty$, then (1.3) holds.

Proof. As noted in the introduction, if $0 < b_n \uparrow \infty$, then (1.2) implies (1.3). To prove (1.2), let $S_n = \sum_{i=1}^n X_i/b_i$. By Lemma 2.1 with δ replaced by $3 + \delta$, we have that for m > n,

$$E(S_{m} - S_{n})^{2} = \sum_{i=n+1}^{m} \frac{EX_{i}^{2}}{b_{i}^{2}} + 2\sum_{i=n+1}^{m-1} \sum_{j=i+1}^{m} \frac{EX_{i}X_{j}}{b_{i}b_{j}}$$

$$\leq \sum_{i=n+1}^{m} \frac{EX_{i}^{2}}{b_{i}^{2}} + \frac{2D^{2}C_{3+\delta}}{\max\{(\log 2)^{3+\delta}, (\log n)^{3+\delta}\}} \sum_{k=1}^{m-n-1} \frac{\rho_{k}}{k} (1 + \log k)^{4+\delta}$$

$$\leq \sum_{i=n+1}^{\infty} \frac{EX_{i}^{2}}{b_{i}^{2}} + \frac{2D^{2}C_{3+\delta}}{\max\{(\log 2)^{3+\delta}, (\log n)^{3+\delta}\}} \sum_{k=1}^{\infty} \frac{\rho_{k}}{k} (1 + \log k)^{4+\delta} \longrightarrow 0$$
(3.1)

as $n \to \infty$ by (i) and (ii). Here $C_{3+\delta} = 2^{\delta+4} \max\{1, (3+\delta)^{3+\delta}e^{-(3+\delta)}\}$. By the Cauchy convergence criterion, there exists a random variable S such that $E(S_n - S)^2 \to 0$ as $n \to \infty$. It is easy to see that $S_{2^n} \to S$ a.s. by the standard method. It remains to show that

$$\max_{2^n < k < 2^{n+1}} |S_k - S_{2^n}| \longrightarrow 0 \text{ a.s. as } n \longrightarrow \infty.$$
 (3.2)

Using Lemma 2.3, (i), and (ii), we get that

$$\sum_{n=1}^{\infty} P\left(\max_{2^{n} < k \le 2^{n+1}} | S_{k} - S_{2^{n}} | > \epsilon\right)
\leq \frac{1}{\epsilon^{2}} \sum_{n=1}^{\infty} E\left(\max_{2^{n} < k \le 2^{n+1}} | S_{k} - S_{2^{n}} |^{2}\right)
\leq \frac{1}{\epsilon^{2}} \sum_{n=1}^{\infty} \left(\frac{\log 2^{n+1}}{\log 2}\right)^{2} \left\{\sum_{i=2^{n}+1}^{2^{n+1}} \frac{EX_{i}^{2}}{b_{i}^{2}} + \frac{2D^{2}C_{3+\delta}}{\left(\log (2^{n}+1)\right)^{3+\delta}} \sum_{k=1}^{2^{n}-1} \frac{\rho_{k}}{k} (1 + \log k)^{4+\delta}\right\}
\leq \frac{1}{\epsilon^{2} (\log 2)^{2}} \sum_{i=3}^{\infty} \frac{(\log(2i))^{2} EX_{i}^{2}}{b_{i}^{2}} + \frac{2D^{2}C_{3+\delta}}{\epsilon^{2} (\log 2)^{3+\delta}} \sum_{n=1}^{\infty} \frac{(n+1)^{2}}{n^{3+\delta}} \sum_{k=1}^{\infty} \frac{\rho_{k}}{k} (1 + \log k)^{4+\delta} < \infty.$$
(3.3)

Then (3.2) follows by the Borel-Cantelli lemma.

Remark 3.2. Hu et al. [2] proved Theorem 3.1 under (i) and (ii)'.

(ii)' $\sum_{n=1}^{\infty} \rho_n / n^q < \infty$ for some $0 \le q < 1$.

Since condition (ii) of Theorem 3.1 is weaker than (ii), Theorem 3.1 improves the result of Hu et al. [2].

We can now establish the following SLLN if condition (2.1) on $\{b_n\}$ is replaced by the condition $0 < b_n \uparrow \infty$.

Theorem 3.3. Let $\{X_n, n \ge 1\}$ be a sequence of square integrable random variables satisfying (1.1). Let $\{b_n, n \ge 1\}$ be a nondecreasing unbounded sequence of positive numbers. Suppose that the following conditions hold:

(i)
$$\sum_{n=1}^{\infty} (\log n)^2 E X_n^2 / b_n^2 < \infty$$
,

(ii)
$$\sum_{n=1}^{\infty} \rho_n \sum_{i=n+1}^{\infty} (\log i)^2 / b_i^2 < \infty$$
,

(iii)
$$\sum_{n=1}^{\infty} EX_n^2 \sum_{i=n+1}^{\infty} \log i / (ib_i^2) < \infty.$$

Then (1.3) holds.

To prove Theorem 3.3, we need the following lemma which is due to Fazekas and Klesov [10].

Lemma 3.4. Let $\{X_n, n \ge 1\}$ be a sequence of random variables and $\{b_n, n \ge 1\}$ be a nondecreasing unbounded sequence of positive numbers. Let $\{\alpha_n, n \ge 1\}$ be a sequence of nonnegative numbers. Assume that for each $n \ge 1$,

$$E\left(\max_{1\leq i\leq n}\left|\sum_{j=1}^{i}X_{j}\right|^{r}\right)\leq \sum_{i=1}^{n}\alpha_{i}, \quad \text{for some constant } r>0. \tag{3.4}$$

If $\sum_{n=1}^{\infty} \alpha_n/b_n^r < \infty$, then (1.3) holds.

Proof of Theorem 3.3. From Lemma 2.2,

$$E\left(\max_{1 \le i \le n} \left| \sum_{j=1}^{i} X_j \right|^2 \right) \le \left(\frac{\log 2n}{\log 2}\right)^2 \left\{ \sum_{i=1}^{n} EX_i^2 + 2\sum_{i=1}^{n-1} \sum_{j=i+1}^{n} (EX_i X_j)^+ \right\}. \tag{3.5}$$

Define $\alpha_n = (\log 2n/\log 2)^2 A_n - (\log 2(n-1)/\log 2)^2 A_{n-1}$ for $n \ge 1$, where $A_0 = 0$ and $A_n = \sum_{i=1}^n EX_i^2 + 2\sum_{i=1}^{n-1} \sum_{j=i+1}^n (EX_iX_j)^+$ for $n \ge 1$. Then $E(\max_{1 \le i \le n} |\sum_{j=1}^i X_j|^2) \le \sum_{i=1}^n \alpha_i$ and

$$\alpha_{n} = \left(\frac{\log 2n}{\log 2}\right)^{2} (A_{n} - A_{n-1}) + A_{n-1} \left\{ \left(\frac{\log 2n}{\log 2}\right)^{2} - \left(\frac{\log 2(n-1)}{\log 2}\right)^{2} \right\}$$

$$= \left(\frac{\log 2n}{\log 2}\right)^{2} \left\{ EX_{n}^{2} + 2\sum_{i=1}^{n-1} (EX_{i}X_{n})^{+} \right\}$$

$$+ \left\{ \left(\frac{\log 2n}{\log 2}\right)^{2} - \left(\frac{\log 2(n-1)}{\log 2}\right)^{2} \right\} \left\{ \sum_{i=1}^{n-1} EX_{i}^{2} + 2\sum_{i=1}^{n-2} \sum_{i=i+1}^{n-1} (EX_{i}X_{j})^{+} \right\}.$$
(3.6)

By Lemma 3.4, it is enough to show that

$$\sum_{n=1}^{\infty} \frac{(\log 2n)^2 E X_n^2}{b_n^2} < \infty, \tag{3.7}$$

$$\sum_{n=1}^{\infty} \frac{(\log 2n)^2}{b_n^2} \sum_{i=1}^{n-1} (EX_i X_n)^+ < \infty, \tag{3.8}$$

$$\sum_{n=2}^{\infty} \frac{(\log 2n)^2 - (\log 2(n-1))^2}{b_n^2} \sum_{i=1}^{n-1} EX_i^2 < \infty, \tag{3.9}$$

$$\sum_{n=3}^{\infty} \frac{(\log 2n)^2 - (\log 2(n-1))^2}{b_n^2} \sum_{i=1}^{n-2} \sum_{j=i+1}^{n-1} (EX_i X_j)^+ < \infty.$$
 (3.10)

Clearly (3.7) holds by (i). It is easy to see that (3.8)–(3.10) hold, and the detailed proofs are omitted. \Box

The following corollary shows that condition (ii) of Theorem 3.3 can be simplified under the additional condition (2.1) on $\{b_n\}$.

Corollary 3.5. Let $\{X_n, n \ge 1\}$ be a sequence of square integrable random variables satisfying (1.1). Let $\{b_n, n \ge 1\}$ be a nondecreasing unbounded sequence of positive numbers satisfying (2.1). Suppose that the following conditions hold:

- (i) $\sum_{n=1}^{\infty} (\log n)^2 E X_n^2 / b_n^2 < \infty$,
- (ii) $\sum_{n=1}^{\infty} (\log n)^2 \rho_n / n < \infty$,
- (iii) $\sum_{n=1}^{\infty} EX_n^2 \sum_{i=n+1}^{\infty} \log i / (ib_i^2) < \infty$.

Then (1.3) holds.

Proof. By (2.1), we have that

$$\sum_{n=1}^{\infty} \rho_n \sum_{i=n+1}^{\infty} \frac{(\log i)^2}{b_i^2} \le D^2 \sum_{n=1}^{\infty} \rho_n \sum_{i=n+1}^{\infty} \frac{(\log i)^2}{i^2} \le C \sum_{n=1}^{\infty} \rho_n \frac{(\log n)^2}{n}$$
(3.11)

for some constant C > 0. Thus the result follows by Theorem 3.3.

Remark 3.6. Condition (ii) of Corollary 3.5 is weaker than condition (ii) of Theorem 3.1. On the other hand, an additional condition is needed in Corollary 3.5 (namely condition (iii) above).

Using the following lemma, we can omit condition (iii) of Theorem 3.3 if conditions (2.1) and (3.12) on $\{b_n\}$ are satisfied. If $C_1n \leq b_n \leq C_2n^{\alpha}$ for all $n \geq 1$ and some constants $C_1 > 0$, $C_2 > 0$, and $\alpha > 0$, then (2.1) and (3.12) hold.

Lemma 3.7. Let $\{b_n, n \geq 1\}$ be a nondecreasing unbounded sequence of positive numbers satisfying (2.1). If

$$\limsup_{n \to \infty} \frac{\log b_n}{\log n} < \infty, \tag{3.12}$$

then

$$\frac{b_n^2}{\log^2 n} \sum_{i=n}^{\infty} \frac{\log i}{ib_i^2} = O(1). \tag{3.13}$$

Proof. Without loss of generality, we may assume that $i \le b_i$ for all $i \ge 1$.

Let fix n. For each $k \ge 1$, define m_k by $m_k = \min\{i \ge n : b_i \ge kb_n\}$. Then $b_{m_k} \ge kb_n$ and $n = m_1 \le m_2 \le \cdots$. It follows that

$$\frac{b_{n}^{2}}{\log^{2}n} \sum_{i=n}^{\infty} \frac{\log i}{ib_{i}^{2}} = \frac{b_{n}^{2}}{\log^{2}n} \sum_{k=1}^{\infty} \sum_{i=m_{k}}^{m_{k+1}-1} \frac{\log i}{ib_{i}^{2}} \le \frac{b_{n}^{2}}{\log^{2}n} \sum_{k=1}^{\infty} \frac{1}{b_{m_{k}}^{2}} \sum_{i=m_{k}}^{m_{k+1}-1} \frac{\log i}{i} \qquad (by \ 0 < b_{n} \uparrow)$$

$$\le \frac{b_{n}^{2}}{\log^{2}n} \sum_{k=1}^{\infty} \frac{1}{(kb_{n})^{2}} \sum_{i=m_{k}}^{m_{k+1}-1} \frac{\log i}{i}$$

$$\le \frac{1}{\log^{2}n} \sum_{k=1}^{\infty} \frac{1}{k^{2}} \left(\sum_{i=1}^{3} \frac{\log i}{i} + \int_{3}^{m_{k+1}-1} \frac{\log x}{x} dx \right)$$

$$\le \frac{1}{\log^{2}n} \sum_{k=1}^{\infty} \frac{1}{k^{2}} \left(\frac{\log 2}{2} + \frac{\log 3}{3} + \frac{(\log(m_{k+1}-1))^{2}}{2} \right),$$
(3.14)

where we assume in the case $m_{k+1} = m_k$, the sum $\sum_{i=m_k}^{m_{k+1}-1} = 0$. Since $b_i \ge i$ for all $i \ge 1$, $b_{\lfloor kb_n \rfloor + 1} \ge \lfloor kb_n \rfloor + 1 \ge kb_n$ and $m_k \le \lfloor kb_n \rfloor + 1 \le kb_n + 1$. So we have that

$$(\log(m_{k+1} - 1))^{2} \le (\log((k+1)b_{n}))^{2} \le 2\{(\log(k+1))^{2} + (\log b_{n})^{2}\}.$$
(3.15)

Substituting this into (3.14), (3.13) holds by (3.12).

The following example shows that Lemma 3.7 fails if (3.12) does not hold.

Example 3.8. Let $\phi(0) = 1$ and $\phi(n) = 2^{\phi(n-1)}$ for $n \ge 1$. Define a sequence $\{b_n, n \ge 1\}$ by $b_n = \phi(k+1)$ if $\phi(k) \le n < \phi(k+1)$. Then $0 < b_n \uparrow \infty$ and $b_n \ge n$ for all $n \ge 1$. Since $b_{\phi(n)} = \phi(n+1)$, we obtain that

$$\frac{\log b_{\phi(n)}}{\log \phi(n)} = \frac{\log \phi(n+1)}{\log \phi(n)} = \frac{\phi(n) \log 2}{\log \phi(n)} \longrightarrow \infty$$
(3.16)

as $n \to \infty$. Hence (3.12) does not hold. We also obtain that for $n \ge 2$,

$$\frac{b_{\phi(n)}^2}{\log^2 \phi(n)} \sum_{i=\phi(n)}^{\infty} \frac{\log i}{i b_i^2} \ge \frac{b_{\phi(n)}^2}{\log^2 \phi(n)} \sum_{i=\phi(n)}^{\phi(n+1)-1} \frac{\log i}{i b_i^2}$$

$$= \frac{1}{\log^2 \phi(n)} \sum_{i=\phi(n)}^{\phi(n+1)-1} \frac{\log i}{i}$$

$$\ge \frac{1}{\log^2 \phi(n)} \int_{\phi(n)}^{\phi(n+1)} \frac{\log x}{x} dx$$

$$= \frac{1}{2} \left(\frac{\phi(n) \log 2}{\log \phi(n)}\right)^2 - \frac{1}{2} \longrightarrow \infty$$
(3.17)

as $n \to \infty$. So (3.13) does not hold.

If $\{b_n\}$ satisfies (2.1) and (3.12), then we can obtain the following SLLN.

Theorem 3.9. Let $\{X_n, n \ge 1\}$ be a sequence of square integrable random variables satisfying (1.1). Let $\{b_n, n \ge 1\}$ be a nondecreasing unbounded sequence of positive numbers satisfying (2.1) and (3.12). Suppose that the following conditions hold:

- (i) $\sum_{n=1}^{\infty} (\log n)^2 E X_n^2 / b_n^2 < \infty$.
- (ii) $\sum_{n=1}^{\infty} (\log n)^2 \rho_n / n < \infty$.

Then (1.3) holds.

Proof. By Lemma 3.7 and (i), we get

$$\sum_{n=1}^{\infty} EX_n^2 \sum_{i=n+1}^{\infty} \frac{\log i}{(ib_i^2)} \le O(1) \sum_{n=1}^{\infty} \frac{\log^2(n+1)EX_n^2}{b_{n+1}^2} < \infty, \tag{3.18}$$

since $\log(n+1)/b_{n+1} \le \log(n+1)/b_n \le 2\log n/b_n$ if $n \ge 2$. The result follows by Corollary 3.5.

Remark 3.10. Under condition (3.12), Theorem 3.9 improves Theorem 3.1, since condition (ii) of Theorem 3.9 is weaker than condition (ii) of Theorem 3.1.

If $b_n = n$ for all $n \ge 1$, then $\{b_n\}$ satisfies (2.1) and (3.12). Hence we can obtain the following.

Corollary 3.11. Let $\{X_n, n \ge 1\}$ be a sequence of square integrable random variables satisfying (1.1). Suppose that the following conditions hold.

- (i) $\sum_{n=1}^{\infty} (\log n)^2 E X_n^2 / n^2 < \infty$.
- (ii) $\sum_{n=1}^{\infty} (\log n)^2 \rho_n / n < \infty$.

Then the SLLN holds. Namely,

$$\frac{\sum_{i=1}^{n} X_i}{n} \longrightarrow 0 \ a.s. \tag{3.19}$$

Remark 3.12. Lyons [3] proved an SLLN (3.19) under the conditions that $EX_n^2 = O(1)$ and

$$\sum_{n=1}^{\infty} \frac{\rho_n}{n} < \infty. \tag{3.20}$$

When $EX_n^2 = O(1)$, condition (i) of Corollary 3.11 is obviously satisfied. Hu et al. [1] proved an SLLN (3.19) under conditions (3.21) and (3.22):

$$\sum_{n=1}^{\infty} \frac{H(n^{\varphi+1})}{n^2} < \infty , \qquad (3.21)$$

$$\sum_{n=1}^{\infty} \frac{\rho_n}{n^{\varphi - 1}} < \infty, \tag{3.22}$$

where $\varphi = (1 + \sqrt{5})/2 (= 1.618 \cdots)$ is the golden ratio, and H(x) > 0 is a nondecreasing function on $(0, \infty)$ such that $EX_n^2 \le H(n)$ for all $n \ge 1$. Condition (ii) of Corollary 3.11 is weaker than (3.22). In general, condition (i) of Corollary 3.11 is not comparable with (3.21).

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