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# Complete convergence and complete moment convergence for weighted sums of extended negatively dependent random variables under sub-linear expectation

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## Abstract

In this paper, we study the complete convergence and complete moment convergence for weighted sums of extended negatively dependent (END) random variables under sub-linear expectations space with the condition of  $C_{\Psi}[|X|^p I(|X|^{1/\alpha})] < \infty$ , further  $\hat{\mathbb{E}}(|X|^p I(|X|^{1/\alpha})) \leq C_{\Psi}[|X|^p I(|X|^{1/\alpha})] < \infty$ ,  $1 < p < 2$  ( $I(x) > 0$  is a slow varying and monotone nondecreasing function). As an application, the Baum-Katz type result for weighted sums of extended negatively dependent random variables is established under sub-linear expectations space. The results obtained in the article are the extensions of the complete convergence and complete moment convergence under classical linear expectation space.

**MSC:** 60F15

**Keywords:** sub-linear expectation space; END random variables; complete convergence; complete moment convergence

## 1 Introduction

Additivity has been generally regarded as a fairly natural assumption, so the classical probability theorems have always been considered under additive probabilities and the linear expectations. However, many uncertain phenomena do not satisfy this assumption. So Peng [1–5] introduced the notions of sub-linear expectations to extend the classical linear expectations. He also established the general theoretical framework of the sub-linear expectation space. The theorems of sub-linear expectations are widely used to assess financial riskiness under uncertainty. For complete convergence and complete moment convergence, there are few reports under sub-linear expectations. This paper aims to obtain the complete convergence and complete moment convergence under sub-linear expectation space with the condition of  $C_{\Psi}[|X|^p I(|X|^{1/\alpha})] < \infty$ , further  $\hat{\mathbb{E}}(|X|^p I(|X|^{1/\alpha})) \leq C_{\Psi}[|X|^p I(|X|^{1/\alpha})] < \infty$ ,  $1 < p < 2$ . In addition, the results and conditions of this paper include a slow varying and monotone nondecreasing function, so the theorems are more generic than the traditional complete convergence. In a word, it is meaningful that this paper extends the complete convergence and complete moment convergence under sub-linear expectation.

Sub-linear expectations generate lots of interesting properties which are unlike those in linear expectations, and the issues in sub-linear expectations are more challenging, so lots of scholars have attached importance to them. Numbers of results have been established, for example, Peng [1–5] gained a weak law of large numbers and a central limit theorem under sub-linear expectation space. Chen [6] gained the law of large numbers for independent identically distributed random variables with the condition of  $\hat{\mathbb{E}}(|X|^{1+\alpha}) < \infty$ . The powerful tools as the moment inequalities and Kolmogorov’s exponential inequalities were established by Zhang [7–9]. He also obtained the Hartman-Wintner’s law of iterated logarithm and Kolmogorov’s strong law of large numbers for identically distributed and extended negatively dependent random variables. Wu and Chen [10] also researched the law of the iterated logarithm, and Cheng [11] studied the strong law of larger number with a general moment condition  $\sup_{i \geq 1} \hat{\mathbb{E}}[|X_i| \psi(|X_i|)] < \infty$ , and so on. Many powerful inequations and conventional methods for linear expectation and probabilities are no longer valid, the study of limit theorems under sub-linear expectation becomes much more challenging.

The complete convergence has a relatively complete development in probability limit theory. The notion of complete convergence was raised by Hsu and Robbins [12], and Chow [13] established complete moment convergence. The complete moment convergence is a more general version of the complete convergence. Lots of results on complete convergence and complete moment convergence for different sequences have been found under classical probability space. For example, Shen *et al.* [14], Wang *et al.* [15] and Wu and Jiang [16], and so on. Some recent papers had new results about complete convergence and complete moment convergence. For instance, Wang *et al.* [17] gained general results of complete convergence and complete moment convergence for weighted sums of some class of random variables, and Wang *et al.* [18] researched complete convergence and complete moment convergence for a class of random variables, and so on. In addition, the theorems of this paper are the extensions of the literature [14] under sub-linear expectation space. And we prove the theorems in this paper with the condition of  $C_V[|X|^p I(|X|^{1/\alpha})] < \infty$ , further  $\hat{\mathbb{E}}(|X|^p I(|X|^{1/\alpha})) \leq C_V[|X|^p I(|X|^{1/\alpha})] < \infty$ ,  $1 < p < 2$  ( $I(x) > 0$  is a slow varying function).

In the next section, we generally introduce some basic notations and concepts, related properties under sub-linear expectations and preliminary lemmas that are useful to prove the main theorems. In Section 3, the complete convergence and complete moment convergence under sub-linear expectation space are established. The proofs of these theorems are stated in the last section.

## 2 Basic settings

The study of this paper uses the framework and notations which are established by Peng [1–5]. So, we omit the definitions of sub-linear expectation ( $\hat{\mathbb{E}}$ ), capacity ( $\mathbb{V}, \nu$ ), countably sub-additive and Choquet integrals/expectations ( $C_V, C_\nu$ ) and so on.

**Definition 2.1** (Peng [1, 2], Zhang [8])

- (i) (Identical distribution) Assume that a space  $\mathbf{X}_1$  and a space  $\mathbf{X}_2$  are two  $n$ -dimensional random vectors defined severally in the sub-linear expectation space  $(\Omega_1, \mathcal{H}_1, \hat{\mathbb{E}}_1)$  and  $(\Omega_2, \mathcal{H}_2, \hat{\mathbb{E}}_2)$ . They are named identically distributed if

$$\hat{\mathbb{E}}_1[\varphi(\mathbf{X}_1)] = \hat{\mathbb{E}}_2[\varphi(\mathbf{X}_2)], \quad \forall \varphi \in C_{l,\text{Lip}}(\mathbb{R}_n),$$

whenever the sub-expectations are finite. A sequence  $\{X_n, n \geq 1\}$  of random variables is named to be identically distributed if, for each  $i \geq 1, X_i$  and  $X_1$  are identically distributed.

- (ii) (Extended negatively dependent) A sequence of random variables  $\{X_n, n \geq 1\}$  is named to be upper (resp. lower) extended negatively dependent if there is some dominating constant  $K \geq 1$  such that

$$\hat{\mathbb{E}}\left(\prod_{i=1}^n g_i(X_i)\right) \leq K \prod_{i=1}^n \hat{\mathbb{E}}(g_i(X_i)), \quad \forall n \geq 2.$$

Whenever the nonnegative functions  $g_i(X_i) \in C_{b,\text{Lip}}(\mathbb{R}), i = 1, 2, \dots,$  are all nondecreasing (resp. all nonincreasing). They are named extended negatively dependent if they are both upper extended negatively dependent and lower extended negatively dependent.

It is distinct that if  $\{X_n, n \geq 1\}$  is a sequence of extended independent random variables and  $f_1(x), f_2(x), \dots \in C_{l,\text{Lip}}(\mathbb{R}),$  then  $\{f_n(X_n), n \geq 1\}$  is also a sequence of extended dependent random variables with  $K = 1;$  if  $\{X_n, n \geq 1\}$  is a sequence of upper extended negatively dependent random variables and  $f_1(x), f_2(x), \dots \in C_{l,\text{Lip}}(\mathbb{R})$  are all nondecreasing (resp. all nonincreasing) functions, then  $\{f_n(X_n); n \geq 1\}$  is also a sequence of upper (resp. lower) extended negatively dependent random variables. It shall be noted that the extended negative dependence of  $\{X_n, n \geq 1\}$  under  $\hat{\mathbb{E}}$  does not imply the extended negative dependence under  $\hat{\mathbb{e}}$ .

In the following, let  $\{X_n, n \geq 1\}$  be a sequence of random variables in  $(\Omega, \mathcal{H}, \hat{\mathbb{E}})$  and  $\sum_{i=1}^n X_i = S_n.$  The symbol  $C$  is on behalf of a generic positive constant which may differ from one place to another. Let  $a_n \ll b_n$  denote that there exists a constant  $C > 0$  such that  $a_n \leq Cb_n$  for sufficiently large  $n, I(\cdot)$  denotes an indicator function,  $a_x \sim b_x$  denotes  $\lim_{x \rightarrow \infty} \frac{a_x}{b_x} = 1.$  Also, let  $a_n \approx b_n$  denote that there exist constants  $c_1 > 0$  and  $c_2 > 0$  such that  $c_1 a_n \leq b_n \leq c_2 a_n$  for sufficiently large  $n.$

The following three lemmas are needed in the proofs of our theorems.

**Lemma 2.1** ([19])  *$l(x)$  is a slow varying function if and only if*

$$l(x) = c(x) \exp\left\{\int_1^x \frac{f(u)}{u} du\right\}, \quad x > 0, \tag{2.1}$$

where  $c(x) \geq 0, \lim_{x \rightarrow \infty} c(x) = c > 0,$  and  $\lim_{x \rightarrow \infty} f(x) = 0.$

**Lemma 2.2** *Suppose  $X \in \mathcal{H}, p > 0, \alpha > 0,$  and  $l(x)$  is a slow varying function.*

- (i) *Then, for  $\forall c > 0,$*

$$C_{\mathbb{V}}[|X|^p l(|X|^{1/\alpha})] < \infty \iff \sum_{n=1}^{\infty} n^{\alpha p - 1} l(n) \mathbb{V}(|X| > cn^{\alpha}) < \infty. \tag{2.2}$$

- (ii) *If  $C_{\mathbb{V}}[|X|^p l(|X|^{1/\alpha})] < \infty,$  then for any  $\theta > 1$  and  $c > 0,$*

$$\sum_{k=1}^{\infty} \theta^{k\alpha p} l(\theta^k) \mathbb{V}(|X| > c\theta^{k\alpha}) < \infty. \tag{2.3}$$

*Proof* (i) By Lemma 2.1, we can express  $l(x)$  as equality (2.1), and  $f(u) \rightarrow 0$  as  $u \rightarrow \infty$ ,  $c(x) \rightarrow c$  as  $x \rightarrow \infty$ . Let  $Z(x) = |x|^p l(|x|^{1/\alpha})$ ,  $Z^{-1}(x)$  be the inverse function of  $Z(x)$ ,  $l(x)$  is a slow varying function and for any  $c > 0$ , we have

$$\begin{aligned} C_V[|X|^p l(|X|^{1/\alpha})] &= \int_0^\infty \mathbb{V}(|X|^p l(|X|^{1/\alpha}) > x) \, dx \\ &= \int_0^\infty \mathbb{V}(|X| > Z^{-1}(x) := cy^\alpha) \, dx \\ &= \int_0^\infty \mathbb{V}(|X| > cy^\alpha) (\alpha p y^{\alpha p-1} l(cy) + y^{\alpha p-1} l(cy) c f(y)) \, dy \\ &\sim \int_0^\infty \mathbb{V}(|X| > cy^\alpha) \alpha p y^{\alpha p-1} l(y) \, dy. \end{aligned}$$

So,

$$C_V[|X|^p l(|X|^{1/\alpha})] < \infty \iff \sum_{n=1}^\infty n^{\alpha p-1} l(n) \mathbb{V}(|X| > cn^\alpha) < \infty.$$

(ii) By the proof of (i), we can imply that for any  $\theta > 1$

$$\begin{aligned} \infty &> \sum_{n=1}^\infty n^{\alpha p-1} l(n) \mathbb{V}(|X| > Cn^\alpha) \\ &\geq C \sum_{k=1}^\infty \sum_{\theta^{k-1} \leq n < \theta^k} \theta^{k(\alpha p-1)} l(\theta^k) \mathbb{V}(|X| > C\theta^{k\alpha}) \\ &\approx \sum_{k=1}^\infty \theta^{k\alpha p} l(\theta^k) \mathbb{V}(|X| > C\theta^{k\alpha}). \end{aligned} \quad \square$$

**Lemma 2.3** (Zhang [9] (Rosenthal’s inequalities)) *Let  $\{X_n, n \geq 1\}$  be a sequence of upper extended negatively dependent random variables in  $(\Omega, \mathcal{H}, \hat{\mathbb{E}})$ . And  $\hat{\mathbb{E}}[X_k] \leq 0, k = 1, \dots, n$ . Then*

$$\mathbb{V}(S_n \geq x) \leq (1 + Ke) \frac{\sum_{k=1}^n \hat{\mathbb{E}}(X_k)^2}{x^2}, \quad \forall x \geq 0. \tag{2.4}$$

**3 Main results**

**Theorem 3.1** *Let  $0 < p < 2, \alpha > 0, \alpha p > 1$ , and  $\{X_n, n \geq 1\}$  be a sequence of END and identically distributed random variables under sub-linear expectations. Let  $l(x) > 0$  be a slow varying and monotone nondecreasing function. And  $\{a_{ni}, 1 \leq i \leq n, n \geq 1\}$  is an array of real numbers such that*

$$\sum_{i=1}^n a_{ni}^2 = O(n). \tag{3.1}$$

If

$$C_V[|X|^p l(|X|^{1/\alpha})] < \infty, \tag{3.2}$$

further, for  $1 < p < 2$ ,

$$\hat{\mathbb{E}}(|X|^p l(|X|^{1/\alpha})) \leq C_V [|X|^p l(|X|^{1/\alpha})]. \tag{3.3}$$

Then, for any  $\varepsilon > 0$ ,

$$\sum_{n=1}^{\infty} n^{\alpha p - 2} l(n) \mathbb{V} \left( \sum_{i=1}^n a_{ni} (X_i - b_i) > \varepsilon n^\alpha \right) < \infty, \tag{3.4}$$

where  $b_i = 0$  if  $p \leq 1$ , and  $b_i = \hat{\mathbb{E}}X_i$  if  $p > 1$ ;

$$\sum_{n=1}^{\infty} n^{\alpha p - 2} l(n) \mathbb{V} \left( \sum_{i=1}^n a_{ni} (X_i - b_i) < -\varepsilon n^\alpha \right) < \infty, \tag{3.5}$$

where  $b_i = 0$  if  $p \leq 1$ , and  $b_i = \hat{\varepsilon}X_i$  if  $p > 1$ .

In particular, if  $\hat{\mathbb{E}}X_i = \hat{\varepsilon}X_i$ , then

$$\sum_{n=1}^{\infty} n^{\alpha p - 2} l(n) \mathbb{V} \left( \left| \sum_{i=1}^n a_{ni} (X_i - b_i) \right| > \varepsilon n^\alpha \right) < \infty, \tag{3.6}$$

where  $b_i = 0$  if  $p \leq 1$ , and  $b_i = \hat{\mathbb{E}}X_i = \hat{\varepsilon}X_i$  if  $p > 1$ .

**Theorem 3.2** Suppose that the conditions of Theorem 3.1 hold, and  $\hat{\mathbb{E}}X_i = \hat{\varepsilon}X_i = b_i$ ,  $1 < p < 2$ , then, for any  $\varepsilon > 0$ ,

$$\sum_{n=1}^{\infty} n^{\alpha p - 2 - \alpha} l(n) C_V \left[ \left| \sum_{i=1}^n a_{ni} (X_i - b_i) \right| - \varepsilon n^\alpha \right]^+ < \infty. \tag{3.7}$$

**Theorem 3.3** Suppose that  $1/2 < \alpha \leq 1$  and other conditions of Theorem 3.1 hold. Let  $l(x) > 0$  be a monotone nondecreasing function. Assume further that  $\{a_{ni}, 1 \leq i \leq n, n \geq 1\}$  is an array of real numbers such that (3.1) holds and  $\hat{\mathbb{E}}X_i = \hat{\varepsilon}X_i = b_i$ . If

$$\hat{\mathbb{E}}(|X|^{1/\alpha} l(|X|^{1/\alpha})) \leq C_V [|X|^{1/\alpha} l(|X|^{1/\alpha})] < \infty, \tag{3.8}$$

then, for  $\forall \varepsilon > 0$ ,

$$\sum_{n=1}^{\infty} \frac{l(n)}{n} \mathbb{V} \left( \sum_{i=1}^n a_{ni} (X_i - b_i) > \varepsilon n^\alpha \right) < \infty. \tag{3.9}$$

**4 Proof**

*Proof of Theorem 3.1* Without loss of generality, we can assume that  $\hat{\mathbb{E}}X_i = 0$ , when  $p > 1$ . We just need to prove (3.5). Because of considering  $\{-X_n; n \geq 1\}$  instead of  $\{X_n; n \geq 1\}$  in (3.5), we can obtain (3.6). Noting that  $a_{ni} \geq 0$ , without loss of generality, we can assume that

$$\sum_{i=1}^n a_{ni}^2 \leq Cn, \tag{4.1}$$

and  $a_{ni} \geq 0$  for all  $1 \leq i \leq n$  and  $n \geq 1$ . It follows by (3.2) and Hölder’s inequality that

$$\sum_{i=1}^n a_{ni} \leq \left( n \sum_{i=1}^n a_{ni}^2 \right)^{1/2} \leq Cn. \tag{4.2}$$

For fixed  $n \geq 1$ , denote for  $1 \leq i \leq n$  that

$$\begin{aligned} X_i^{(n)} &= -n^\alpha I(X_i < -n^\alpha) + X_i I(|X_i| \leq n^\alpha) + n^\alpha I(X_i > n^\alpha), \\ T^{(n)} &= n^{-\alpha} \sum_{i=1}^k a_{ni} (X_i^{(n)} - \hat{\mathbb{E}}X_i^{(n)}). \end{aligned}$$

It is easily checked that for  $\forall \varepsilon > 0$ ,

$$\left( \sum_{i=1}^n a_{ni} X_i > \varepsilon n^\alpha \right) \subset \bigcup_{i=1}^n (|X_i| > n^\alpha) \cup \left( \sum_{i=1}^n a_{ni} X_i^{(n)} > \varepsilon n^\alpha \right),$$

which can imply that

$$\begin{aligned} &\sum_{n=1}^\infty n^{\alpha p-2} l(n) \mathbb{V} \left( \sum_{i=1}^n a_{ni} X_i > \varepsilon n^\alpha \right) \\ &\leq \sum_{n=1}^\infty n^{\alpha p-2} l(n) \sum_{i=1}^n \mathbb{V}(|X_i| > n^\alpha) \\ &\quad + \sum_{n=1}^\infty n^{\alpha p-2} l(n) \mathbb{V} \left( T^{(n)} > \varepsilon - \left| n^{-\alpha} \sum_{i=1}^n a_{ni} \hat{\mathbb{E}}X_i^{(n)} \right| \right) \\ &:= I_1 + I_2. \end{aligned}$$

For  $0 < \mu < 1$ , let  $g(x)$  be a decreasing function and  $g(x) \in C_{l,\text{Lip}}(\mathbb{R})$ ,  $0 \leq g(x) \leq 1$  for all  $x$  and  $g(x) = 1$  if  $|x| \leq \mu$ ,  $g(x) = 0$  if  $|x| > 1$ . Then

$$I(|x| \leq \mu) \leq g(x) \leq I(|x| \leq 1), \quad I(|x| \geq 1) \leq 1 - g(x) \leq I(|x| \geq \mu). \tag{4.3}$$

In order to prove (3.5), it suffices to show  $I_1 < \infty$  and  $I_2 < \infty$ . By Lemma 2.2(i) and identically distributed random variables, we can get that

$$\begin{aligned} I_1 &\leq C \sum_{n=1}^\infty n^{\alpha p-2} l(n) \sum_{i=1}^n \hat{\mathbb{E}} \left( 1 - g \left( \frac{X_i}{n^\alpha} \right) \right) \\ &\leq C \sum_{n=1}^\infty n^{\alpha p-1} l(n) \hat{\mathbb{E}} \left( 1 - g \left( \frac{X}{n^\alpha} \right) \right) \\ &\leq C \sum_{n=1}^\infty n^{\alpha p-1} l(n) \mathbb{V}(|X| > \mu n^\alpha) < \infty. \end{aligned}$$

In the following, we prove that  $I_2 < \infty$ . First, we prove that

$$\left| n^{-\alpha} \sum_{i=1}^n a_{ni} \hat{\mathbb{E}}X_i^{(n)} \right| \rightarrow 0 \quad \text{as } n \rightarrow \infty. \tag{4.4}$$

Case 1:  $0 < p \leq 1$ .

For any  $r > 0$ , by the  $C_r$  inequality and (4.4),

$$\begin{aligned} |X^{(n)}|^r &\ll |X|^r I(|X| \leq n^\alpha) + n^{r\alpha} I(|X| > n^\alpha) \leq |X|^r g\left(\frac{\mu X}{n^\alpha}\right) + n^{r\alpha} \left(1 - g\left(\frac{X}{n^\alpha}\right)\right), \\ \widehat{\mathbb{E}}|X^{(n)}|^r &\ll \widehat{\mathbb{E}}\left[|X|^r g\left(\frac{\mu X}{n^\alpha}\right)\right] + n^{r\alpha} \widehat{\mathbb{E}}\left[1 - g\left(\frac{X}{n^\alpha}\right)\right] \\ &\leq \widehat{\mathbb{E}}\left[|X|^r g\left(\frac{\mu X}{n^\alpha}\right)\right] + n^{r\alpha} \mathbb{V}(|X| > \mu n^\alpha). \end{aligned} \tag{4.5}$$

So, by (4.3) we can imply that

$$\begin{aligned} \left| n^{-\alpha} \sum_{i=1}^n a_{ni} \widehat{\mathbb{E}}X_i^{(n)} \right| &\ll n^{-\alpha} \widehat{\mathbb{E}}|X^{(n)}| \sum_{i=1}^n a_{ni} \\ &\leq n^{1-\alpha} \widehat{\mathbb{E}}|X^{(n)}| \\ &\leq Cn^{1-\alpha} \left( \widehat{\mathbb{E}}|X| g\left(\frac{\mu X}{n^\alpha}\right) + n^\alpha \mathbb{V}(|X| > \mu n^\alpha) \right) \\ &\leq Cn^{1-\alpha} \widehat{\mathbb{E}}|X| g\left(\frac{\mu X}{n^\alpha}\right) + Cn \mathbb{V}(|X| > \mu n^\alpha) \\ &:= I_{21} + Cn \mathbb{V}(|X| > \mu n^\alpha). \end{aligned} \tag{4.6}$$

By (2.3), we can imply that

$$\infty > \sum_{n=1}^{\infty} n^{\alpha p - 1} l(n) \mathbb{V}(|X| > cn^\alpha) \geq \sum_{n=1}^{\infty} \mathbb{V}(|X| > cn^\alpha),$$

and  $\mathbb{V}(|X| > \mu n^\alpha) \downarrow$ , so we get  $n \mathbb{V}(|X| > \mu n^\alpha) \rightarrow 0$  as  $n \rightarrow \infty$ . Next, we estimate  $I_{21}$ . Let  $g_j(x) \in C_{l,\text{Lip}}(\mathbb{R})$ ,  $j \geq 1$  such that  $0 \leq g_j(x) \leq 1$  for all  $x$  and  $g_j(\frac{x}{2^{j\alpha}}) = 1$  if  $2^{(j-1)\alpha} < |x| \leq 2^{j\alpha}$ ,  $g_j(\frac{x}{2^{j\alpha}}) = 0$  if  $|x| \leq 2^{(j-1)\alpha}$  or  $|x| > (1 + \mu)2^{j\alpha}$ . Then

$$g_j\left(\frac{X}{2^{j\alpha}}\right) \leq I(\mu 2^{(j-1)\alpha} < |X| \leq (1 + \mu)2^{j\alpha}), \quad X^r g\left(\frac{X}{2^{k\alpha}}\right) \leq 1 + \sum_{j=1}^k X^r g_j\left(\frac{X}{2^{j\alpha}}\right). \tag{4.7}$$

For every  $n$ , there exists  $k$  such that  $2^{k-1} \leq n < 2^k$ , thus by (4.7),  $g(x) \downarrow$ , and  $n^{-\alpha+1} \downarrow 0$ , from  $\alpha > \frac{1}{p} \geq 1$ , we get

$$\begin{aligned} I_{21} &\leq 2^{(k-1)(1-\alpha)} \widehat{\mathbb{E}}|X| g\left(\frac{\mu X}{2^{k\alpha}}\right) \\ &\leq C 2^{(k-1)(1-\alpha)} \sum_{j=1}^k \widehat{\mathbb{E}}|X| g_j\left(\frac{\mu X}{2^{j\alpha}}\right) \\ &\leq 2^{(k-1)(1-\alpha)} \sum_{j=1}^k 2^{j\alpha} \mathbb{V}(|X| > 2^{(j-1)\alpha}). \end{aligned}$$

Noting that by (2.4),  $\alpha p > 1$ ,

$$\begin{aligned} \sum_{j=1}^{\infty} \frac{2^{j\alpha}}{2^{j(\alpha-1)}} \mathbb{V}(|X| > 2^{(j-1)\alpha}) &= \sum_{j=1}^{\infty} 2^j \mathbb{V}(|X| > 2^{-\alpha} 2^{j\alpha}) \\ &\leq \sum_{j=1}^{\infty} 2^{j\alpha p} l(2^j) \mathbb{V}(|X| > 2^{-\alpha} 2^{j\alpha}) < \infty. \end{aligned}$$

It follows that

$$I_{21} \rightarrow 0 \quad \text{as } n \rightarrow \infty$$

from the Kronecker lemma and  $2^{j(\alpha-1)} \uparrow \infty$ .

Case 2:  $1 < p < 2$ .

By (3.4), we can get that

$$\hat{\mathbb{E}}|X|^p < \infty. \tag{4.8}$$

By (4.9) and  $\alpha p > 1, 1 < p < 2$ , one can get that

$$\begin{aligned} \left| n^{-\alpha} \sum_{i=1}^n a_{ni} \hat{\mathbb{E}}X_i^{(n)} \right| &\leq n^{-\alpha} \sum_{i=1}^n a_{ni} |\hat{\mathbb{E}}X_i - \hat{\mathbb{E}}X_i^{(n)}| \\ &\leq n^{-\alpha} \sum_{i=1}^n a_{ni} \hat{\mathbb{E}}|X_i - X_i^{(n)}| \\ &\leq n^{1-\alpha} \frac{\hat{\mathbb{E}}|X| |X|^{p-1}}{n^{\alpha(p-1)}} \left( 1 - g\left(\frac{X}{n^\alpha}\right) \right) \\ &\ll C n^{1-\alpha p} \hat{\mathbb{E}}|X|^p \rightarrow 0 \quad \text{as } n \rightarrow \infty. \end{aligned}$$

It follows that for all  $n$  large enough,

$$\left| n^{-\alpha} \sum_{i=1}^n a_{ni} \hat{\mathbb{E}}X_i^{(n)} \right| < \varepsilon/2,$$

which implies that

$$I_2 \leq C \sum_{n=1}^{\infty} n^{\alpha p-2} l(n) \mathbb{V}(T^{(n)} > \varepsilon/2).$$

By Definition 2.1(ii), we can know that fixed  $n \geq 1, \{a_{ni}(X_i^{(n)} - \hat{\mathbb{E}}X_i^{(n)}), 1 \leq i \leq n\}$  are still END random variables. Hence, we have by Lemma 2.3 (taking  $x = \varepsilon n^\alpha$ ) that

$$\begin{aligned} \mathbb{V}(T^{(n)} > \varepsilon/2) &\leq C \frac{\sum_{i=1}^n \hat{\mathbb{E}}(a_{ni}(X_i^{(n)} - \hat{\mathbb{E}}X_i^{(n)}))^2}{\varepsilon^2 n^{2\alpha}} \\ &\leq C n^{-2\alpha} \sum_{i=1}^n a_{ni}^2 \hat{\mathbb{E}}(X_i^{(n)})^2. \end{aligned}$$



By (4.6), we have

$$\begin{aligned}
 I_2 &\leq C \sum_{n=1}^{\infty} n^{\alpha p - 2\alpha - 2} l(n) \sum_{i=1}^n a_{ni}^2 \hat{\mathbb{E}} |X_i^{(n)}|^2 \\
 &\leq C \sum_{n=1}^{\infty} n^{\alpha p - 2\alpha - 1} l(n) \hat{\mathbb{E}} \left[ X^2 g \left( \frac{\mu X}{n^\alpha} \right) \right] \\
 &\quad + C \sum_{n=1}^{\infty} n^{\alpha p - 1} l(n) \mathbb{V}(|X| > \mu n^\alpha) \\
 &:= I_3 + I_4.
 \end{aligned}$$

By Lemma 2.2(i), we can get  $I_4 < \infty$ . Noting that by (4.8)

$$\begin{aligned}
 I_3 &= C \sum_{j=0}^{\infty} \sum_{n=2^j}^{2^{j+1}-1} n^{\alpha p - 2\alpha - 1} l(n) \hat{\mathbb{E}} \left[ X^2 g \left( \frac{\mu X}{n^\alpha} \right) \right] \\
 &\leq C \sum_{j=1}^{\infty} 2^{(\alpha p - 2\alpha - 1)j} l(2^j) \hat{\mathbb{E}} \left[ X^2 g \left( \frac{\mu X}{2^{\alpha(j+1)}} \right) \right] \\
 &\leq C \sum_{j=1}^{\infty} 2^{\alpha(p-2)j} l(2^j) \hat{\mathbb{E}} \left[ 1 + \sum_{k=1}^j X^2 g_k \left( \frac{\mu X}{2^{\alpha(k+1)}} \right) \right] \\
 &\leq C \sum_{j=1}^{\infty} 2^{\alpha(p-2)j} l(2^j) + C \sum_{j=1}^{\infty} 2^{\alpha(p-2)j} l(2^j) \sum_{k=1}^j \hat{\mathbb{E}} \left[ X^2 g_k \left( \frac{\mu X}{2^{\alpha(k+1)}} \right) \right] \\
 &= I_{31} + I_{32}.
 \end{aligned}$$

By  $p < 2$ , we get  $I_{31} < \infty$ . Next we estimate  $I_{32}$ . By (2.4), we can imply that

$$\begin{aligned}
 I_{32} &= \sum_{j=1}^{\infty} 2^{\alpha(p-2)j} l(2^j) \sum_{k=1}^j \hat{\mathbb{E}} \left[ X^2 g_k \left( \frac{\mu X}{2^{\alpha(k+1)}} \right) \right] \\
 &\leq \sum_{k=1}^{\infty} 2^{\alpha p k} l(2^k) \hat{\mathbb{E}} \left[ g_k \left( \frac{\mu X}{2^{\alpha(k+1)}} \right) \right] \\
 &\leq \sum_{k=1}^{\infty} 2^{\alpha p k} l(2^k) \mathbb{V}(|X| > 2^{\alpha k}) \\
 &< \infty.
 \end{aligned}$$

Hence, it follows that

$$I_3 < \infty.$$

By  $I_3 < \infty$  and  $I_4 < \infty$ , we can get  $I_2 < \infty$ .

This finishes the proof of Theorem 3.1. □

*Proof of Theorem 3.2* Without loss of generality, we can assume that  $\hat{\mathbb{E}}X_i = 0$  when  $p > 1$ , and assume that  $a_{ni} \geq 0$ . For  $\forall \varepsilon > 0$ , we have by Theorem 3.1 that

$$\begin{aligned} & \sum_{n=1}^{\infty} n^{\alpha p-2-\alpha} l(n) C_{\mathbb{V}} \left( \sum_{i=1}^n a_{ni} (X_i - b_i) - \varepsilon n^{\alpha} \right)^+ \\ &= \sum_{n=1}^{\infty} n^{\alpha p-2-\alpha} l(n) \int_0^{\infty} \mathbb{V} \left( \sum_{i=1}^n a_{ni} X_i - \varepsilon n^{\alpha} > t \right) dt \\ &= \sum_{n=1}^{\infty} n^{\alpha p-2-\alpha} l(n) \int_0^{n^{\alpha}} \mathbb{V} \left( \sum_{i=1}^n a_{ni} X_i - \varepsilon n^{\alpha} > t \right) dt \\ &\quad + \sum_{n=1}^{\infty} n^{\alpha p-2-\alpha} l(n) \int_{n^{\alpha}}^{\infty} \mathbb{V} \left( \sum_{i=1}^n a_{ni} X_i - \varepsilon n^{\alpha} > t \right) dt \\ &\leq \sum_{n=1}^{\infty} n^{\alpha p-2} l(n) \mathbb{V} \left( \sum_{i=1}^n a_{ni} X_i > \varepsilon n^{\alpha} \right) \\ &\quad + \sum_{n=1}^{\infty} n^{\alpha p-2-\alpha} l(n) \int_{n^{\alpha}}^{\infty} \mathbb{V} \left( \sum_{i=1}^n a_{ni} X_i - \varepsilon n^{\alpha} > t \right) dt \\ &\leq C \sum_{n=1}^{\infty} n^{\alpha p-2-\alpha} l(n) \int_{n^{\alpha}}^{\infty} \mathbb{V} \left( \sum_{i=1}^n a_{ni} X_i > t \right) dt. \end{aligned}$$

Hence, it suffices to show that

$$H := \sum_{n=1}^{\infty} n^{\alpha p-2-\alpha} l(n) \int_{n^{\alpha}}^{\infty} \mathbb{V} \left( \sum_{i=1}^n a_{ni} X_i > t \right) dt < \infty.$$

For  $t > n^{\alpha}$ , denote

$$Z_{ti} = -tI(X_i < -t) + X_i I(|X_i| \leq t) + tI(X_i > t), \quad i = 1, 2, \dots \tag{4.9}$$

and

$$U_{ti} = tI(X_i < -t) + X_i I(|X_i| > t) - tI(X_i > t), \quad i = 1, 2, \dots \tag{4.10}$$

Since  $X_i = U_{ti} + Z_{ti}$ , it follows that

$$\begin{aligned} H &\leq \sum_{n=1}^{\infty} n^{\alpha p-2-\alpha} l(n) \int_{n^{\alpha}}^{\infty} \mathbb{V} \left( \sum_{i=1}^n a_{ni} X_i > t \right) dt \\ &\leq \sum_{n=1}^{\infty} n^{\alpha p-2-\alpha} l(n) \int_{n^{\alpha}}^{\infty} \mathbb{V} \left( \left| \sum_{i=1}^n a_{ni} U_{ti} \right| > t/2 \right) dt \\ &\quad + \sum_{n=1}^{\infty} n^{\alpha p-2-\alpha} l(n) \int_{n^{\alpha}}^{\infty} \mathbb{V} \left( t^{-1} \sum_{i=1}^n a_{ni} (Z_{ti} - \hat{\mathbb{E}}Z_{ti}) > 1/2 - t^{-1} \left| \sum_{i=1}^n a_{ni} \hat{\mathbb{E}}Z_{ti} \right| \right) dt \\ &:= H_1 + H_2. \end{aligned}$$

Note that by Lemma 2.2(i)

$$\begin{aligned}
 H_1 &\leq C \sum_{n=1}^{\infty} n^{\alpha p-2-\alpha} l(n) \int_{n^\alpha}^{\infty} \mathbb{V}(\exists 1 \leq i < n, |X_i| > t) dt \\
 &\leq C \sum_{n=1}^{\infty} n^{\alpha p-2-\alpha} l(n) \int_{n^\alpha}^{\infty} \sum_{i=1}^n \mathbb{V}(|X_i| > t) dt \\
 &\leq C \sum_{n=1}^{\infty} n^{\alpha p-1-\alpha} l(n) \int_{n^\alpha}^{\infty} \hat{\mathbb{E}}\left(1 - g\left(\frac{X}{t}\right)\right) dt \\
 &= C \sum_{n=1}^{\infty} n^{\alpha p-1-\alpha} l(n) \sum_{m=n}^{\infty} \int_{m^\alpha}^{(m+1)^\alpha} \hat{\mathbb{E}}\left(1 - g\left(\frac{X}{t}\right)\right) dt \\
 &\leq C \sum_{n=1}^{\infty} n^{\alpha p-1-\alpha} l(n) \sum_{m=n}^{\infty} [(m+1)^\alpha - m^\alpha] \hat{\mathbb{E}}\left(1 - g\left(\frac{X}{m^\alpha}\right)\right) \\
 &\leq C \sum_{m=1}^{\infty} m^{\alpha-1} \mathbb{V}(|X| > \mu m^\alpha) \sum_{n=1}^m n^{\alpha p-1-\alpha} l(n) \\
 &\ll \sum_{m=1}^{\infty} m^{\alpha p-1} l(m) \mathbb{V}(|X| > \mu m^\alpha) < \infty.
 \end{aligned}
 \tag{4.11}$$

In the following, we prove that  $H_2 < \infty$ . First, we show that

$$\sup_{t \geq n^\alpha} t^{-1} \left| \sum_{i=1}^n a_{ni} \hat{\mathbb{E}} Z_{ti} \right| \rightarrow 0 \quad \text{as } n \rightarrow \infty.
 \tag{4.12}$$

Case 1:  $0 < p \leq 1$ .

Note (4.10) and (4.4), which imply that

$$\begin{aligned}
 |Z_{ii}| &\ll |X_i| I(|X_i| \leq t) + t I(|X_i| > t) \leq |X_i| g\left(\frac{\mu X_i}{t}\right) + t \left(1 - g\left(\frac{X_i}{t}\right)\right), \\
 \hat{\mathbb{E}}|Z_{ii}| &\ll \hat{\mathbb{E}}\left[|X| g\left(\frac{\mu X}{t}\right)\right] + t \hat{\mathbb{E}}\left[1 - g\left(\frac{X}{t}\right)\right] \\
 &\leq \hat{\mathbb{E}}\left[|X| g\left(\frac{\mu X}{t}\right)\right] + t \mathbb{V}(|X| > \mu t).
 \end{aligned}
 \tag{4.13}$$

So, for  $t > n^\alpha$ , we get

$$\begin{aligned}
 \sup_{t \geq n^\alpha} t^{-1} \left| \sum_{i=1}^n a_{ni} \hat{\mathbb{E}} Z_{ti} \right| &\ll \sup_{t \geq n^\alpha} t^{-1} n \hat{\mathbb{E}}|Z_{ii}| \\
 &\leq \sup_{t \geq n^\alpha} t^{-1} n \left( \hat{\mathbb{E}}|X| g\left(\frac{\mu X}{t}\right) + t \mathbb{V}(|X| > \mu t) \right) \\
 &\leq n^{1-\alpha} \hat{\mathbb{E}}|X| g\left(\frac{\mu X}{n^\alpha}\right) + n \mathbb{V}(|X| > \mu n^\alpha) \\
 &:= H_{21} + n \mathbb{V}(|X| > \mu n^\alpha).
 \end{aligned}$$

We get  $n\mathbb{V}(|X| > \mu n^\alpha) \rightarrow 0$  as  $n \rightarrow \infty$  in the proof of (4.7). Next, we estimate  $H_{21}$ . For every  $n$ , there exists  $k$  such that  $2^{k-1} \leq n < 2^k$ , thus by (4.8), (4.13),  $g(x) \downarrow$ ,  $t > n^\alpha$  and  $n^{-\alpha+1} \downarrow 0$ , from  $\alpha > 1$ , we get

$$\begin{aligned} H_{21} &\leq Cn^{1-\alpha} \hat{\mathbb{E}}|X|g\left(\frac{\mu X}{n^\alpha}\right) \\ &\leq 2^{(k-1)(1-\alpha)} \hat{\mathbb{E}}|X|g\left(\frac{\mu X}{2^{k\alpha}}\right) \\ &\leq 2^{(k-1)(1-\alpha)} \sum_{j=1}^k \hat{\mathbb{E}}|X|g\left(\frac{\mu X}{2^{j\alpha}}\right) \\ &\leq 2^{(k-1)(1-\alpha)} \sum_{j=1}^k 2^{j\alpha} \mathbb{V}(|X| > 2^{(j-1)\alpha}). \end{aligned}$$

Noting that by (2.4),  $\alpha p > 1$ ,

$$\begin{aligned} \sum_{j=1}^{\infty} \frac{2^{j\alpha}}{2^{j(\alpha-1)}} \mathbb{V}(|X| > 2^{(j-1)\alpha}) &= \sum_{j=1}^{\infty} 2^j \mathbb{V}(|X| > 2^{-\alpha} 2^{j\alpha}) \\ &\leq \sum_{j=1}^{\infty} 2^{j\alpha p} l(2^j) \mathbb{V}(|X| > 2^{-\alpha} 2^{j\alpha}) < \infty. \end{aligned}$$

It follows that

$$H_{21} \rightarrow 0 \quad \text{as } n \rightarrow \infty$$

from the Kronecker lemma and  $2^{j(\alpha-1)} \uparrow \infty$ .

Case 2:  $1 < p < 2$ .

By  $\hat{\mathbb{E}}X_i = 0$  and  $\alpha p > 1$ ,  $t > n^\alpha$ , we can get that

$$\begin{aligned} \sup_{t \geq n^\alpha} t^{-1} \left| \sum_{i=1}^n a_{ni} \hat{\mathbb{E}}Z_{ti} \right| &\leq \sup_{t \geq n^\alpha} t^{-1} \sum_{i=1}^n a_{ni} |\hat{\mathbb{E}}X_i - \hat{\mathbb{E}}Z_{ti}| \\ &\leq n^{-\alpha} \sum_{i=1}^n a_{ni} \hat{\mathbb{E}}|X_i - X_i^{(n)}| \\ &\leq Cn^{1-\alpha} \frac{\hat{\mathbb{E}}|X||X|^{p-1}}{n^{\alpha(p-1)}} \left( 1 - g\left(\frac{X}{n^\alpha}\right) \right) \\ &= Cn^{1-\alpha p} \hat{\mathbb{E}}|X|^p \left( 1 - g\left(\frac{X}{n^\alpha}\right) \right) \rightarrow 0 \quad \text{as } n \rightarrow \infty. \end{aligned}$$

It follows that for all  $n$  large enough,

$$t^{-1} \left| \sum_{i=1}^n a_{ni} \hat{\mathbb{E}}Z_{it} \right| < 1/4,$$

which implies that

$$H_2 \leq C \sum_{n=1}^{\infty} n^{\alpha p-2} l(n) \int_{n^\alpha}^{\infty} \mathbb{V} \left( t^{-1} \sum_{i=1}^n a_{ni} (Z_{ti} - \hat{\mathbb{E}} Z_{ti}) > 1/4 \right) dt.$$

For fixed  $t > n^\alpha$  and  $n \geq 1$ , it is easily seen that  $\{a_{ni}(Z_{ti} - \hat{\mathbb{E}} Z_{ti}), i \geq 1\}$  are still END random variables. Hence, we have by Markov’s inequality, Lemma 2.3, (4.3), (4.12), (4.13), Lemma 2.2(i) that

$$\begin{aligned} H_2 &\leq C \sum_{n=1}^{\infty} n^{\alpha p-2-\alpha} l(n) \int_{n^\alpha}^{\infty} t^{-2} \sum_{i=1}^n a_{ni}^2 \hat{\mathbb{E}} Z_{ti}^2 dt \\ &\leq C \sum_{n=1}^{\infty} n^{\alpha p-1-\alpha} l(n) \int_{n^\alpha}^{\infty} t^{-2} \hat{\mathbb{E}} X^2 g \left( \frac{\mu X}{t} \right) dt \\ &\quad + C \sum_{n=1}^{\infty} n^{\alpha p-1-\alpha} l(n) \int_{n^\alpha}^{\infty} \hat{\mathbb{E}} \left( 1 - g \left( \frac{X}{t} \right) \right) dt \\ &\leq C \sum_{n=1}^{\infty} n^{\alpha p-1-\alpha} l(n) \sum_{m=n}^{\infty} \int_{m^\alpha}^{(m+1)^\alpha} t^{-2} \hat{\mathbb{E}} X^2 g \left( \frac{\mu X}{t} \right) dt \\ &\leq C \sum_{n=1}^{\infty} n^{\alpha p-1-\alpha} l(n) \sum_{m=n}^{\infty} m^{\alpha-1-2\alpha} \hat{\mathbb{E}} X^2 g \left( \frac{\mu X}{(m+1)^\alpha} \right) dt \\ &= C \sum_{m=1}^{\infty} m^{\alpha-1-2\alpha} \hat{\mathbb{E}} X^2 g \left( \frac{\mu X}{(m+1)^\alpha} \right) \sum_{n=1}^m n^{\alpha p-1-\alpha} l(n) \\ &\leq C \sum_{m=1}^{\infty} m^{\alpha-1-2\alpha} \hat{\mathbb{E}} X^2 g \left( \frac{\mu X}{(m+1)^\alpha} \right) m^{\alpha p-\alpha} l(m) \\ &= C \sum_{n=1}^{\infty} n^{\alpha p-1-2\alpha} l(n) \hat{\mathbb{E}} X^2 g \left( \frac{\mu X}{(n+1)^\alpha} \right) \\ &\leq C \sum_{n=1}^{\infty} n^{\alpha p-1} l(n) \mathbb{V}(|X| > \mu n^\alpha) < \infty. \end{aligned}$$

Hence, this finishes the proof of Theorem 3.2. □

*Proof of Theorem 3.3* We use the same notations as those in Theorem 3.1. The proof is similar to that of Theorem 3.1. We only need to show that

$$n^{-\alpha} \left| \sum_{i=1}^n \hat{\mathbb{E}} a_{ni} X_i^{(n)} \right| \rightarrow 0 \quad \text{as } n \rightarrow \infty.$$

Because  $l(x) > 0$  is a monotone nondecreasing function, we have

$$\begin{aligned} |X|^{1/\alpha} &= |X|^{1/\alpha} I(|X| \leq 1) + |X|^{1/\alpha} I(|X|^{1/\alpha}) \frac{1}{l(|X|^{1/\alpha})} I(|X| > 1) \\ &\leq 1 + |X|^{1/\alpha} l(|X|^{1/\alpha}) \frac{1}{l(1)}, \end{aligned}$$

which together with (3.8) yields that  $C_{\mathbb{V}}|X|^{1/\alpha} < C_{\mathbb{V}}[|X|^{1/\alpha}I(|X|^{1/\alpha})] < \infty$ . Noting that  $1 \leq 1/\alpha < 2$  and  $\hat{\mathbb{E}}X_i = 0$ , we have

$$\begin{aligned} n^{-\alpha} \left| \sum_{i=1}^n \hat{\mathbb{E}}a_{ni}X_i^{(n)} \right| &\leq n^{-\alpha} \sum_{i=1}^n a_{ni} |\hat{\mathbb{E}}X_i - \hat{\mathbb{E}}X_i^{(n)}| \\ &\leq n^{-\alpha} \sum_{i=1}^n a_{ni} \hat{\mathbb{E}}|X_i - X_i^{(n)}| \\ &\leq Cn^{1-\alpha} \hat{\mathbb{E}}|X| \left( 1 - g\left(\frac{X}{n^\alpha}\right) \right) \\ &\leq Cn^{1-\alpha} \frac{\hat{\mathbb{E}}|X||X|^{1/\alpha-1}}{n^{1-\alpha}} \left( 1 - g\left(\frac{X}{n^\alpha}\right) \right) \\ &\ll C_{\mathbb{V}}(|X|^{1/\alpha}I(|X| > \mu n^\alpha)) \rightarrow 0 \quad \text{as } n \rightarrow \infty. \quad \square \end{aligned}$$

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**Authors' contributions**

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