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# Complete moment convergence for moving average process generated by $\rho^-$ -mixing random variables

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## Abstract

Let  $\{Y_i, -\infty < i < \infty\}$  be a sequence of  $\rho^-$ -mixing random variables without the assumption of identical distributions, and  $\{a_i, -\infty < i < \infty\}$  be an absolutely summable sequence of real numbers. In this paper, under some suitable conditions, we establish the complete moment convergence for the partial sum of moving average processes  $\{X_n = \sum_{i=-\infty}^{\infty} a_i Y_{i+n}, n \geq 1\}$ . These results promote and improve the corresponding results obtained by Li and Zhang (Stat. Probab. Lett. 70:191-197, 2004) from NA to the case of a  $\rho^-$ -mixing setting.

**Keywords:** complete moment convergence; moving average process;  $\rho^-$ -mixing; Marcinkiewicz-Zygmund strong law of large numbers

## 1 Introduction

Let  $\{Y_i, -\infty < i < \infty\}$  be a sequence of random variables and  $\{a_i, -\infty < i < \infty\}$  be an absolutely summable sequence of real numbers, and for  $n \geq 1$  set  $X_n = \sum_{i=-\infty}^{\infty} a_i Y_{i+n}$ . The limit behavior of the moving average process  $\{X_n, n \geq 1\}$  has been extensively investigated by many authors. For example, Baek *et al.* [1] have obtained the convergence of moving average processes, Burton and Dehling [2] have obtained a large deviation principle, Ibragimov [3] has established the central limit theorem, Račkauskas and Suquet [4] have proved the functional central limit theorems for self-normalized partial sums of linear processes, and Chen *et al.* [5], Guo [6], Kim *et al.* [7, 8], Ko *et al.* [9], Li *et al.* [10], Li and Zhang [11], Qiu *et al.* [12], Wang and Hu [13], Yang and Hu [14], Zhang [15], Zhen *et al.* [16], Zhou *et al.* [17], Zhou and Lin [18], Shen *et al.* [19] have obtained the complete (moment) convergence of moving average process based on a sequence of dependent (or mixing) random variables, respectively. But very few results for moving average process based on a  $\rho^-$ -mixing random variables are known. Firstly, we recall some definitions.

For two nonempty disjoint sets  $S, T$  of real numbers, we define  $\text{dist}(S, T) = \min\{|j - k|; j \in S, k \in T\}$ . Let  $\sigma(S)$  be the  $\sigma$ -field generated by  $\{Y_k, k \in S\}$ , and define  $\sigma(T)$  similarly.

**Definition 1.1** A sequence  $\{Y_i, -\infty < i < \infty\}$  is called  $\rho^-$ -mixing, if

$$\rho^-(s) = \sup\{\rho^-(S, T); S, T \subset \mathbb{Z}, \text{dist}(S, T) \geq s\} \rightarrow 0 \quad \text{as } s \rightarrow \infty,$$

where

$$\rho^-(S, T) = 0 \vee \sup\{\text{corr}(f(X_i, i \in S), g(X_j, j \in T)),$$

where the supremum is taken over all coordinatewise increasing real functions  $f$  on  $R^S$  and  $g$  on  $R^T$ .

**Definition 1.2** A sequence  $\{Y_i, -\infty < i < \infty\}$  is called  $\rho^*$ -mixing if

$$\rho^*(s) = \sup\{\rho(S, T); S, T \subset Z, \text{dist}(S, T) \geq s\} \rightarrow 0 \quad \text{as } s \rightarrow \infty,$$

where

$$\rho(S, T) = \sup\{|\text{corr}(f, g)|; f \in L_2(\sigma(S)), g \in L_2(\sigma(T))\}.$$

**Definition 1.3** A sequence  $\{Y_i, i \in Z\}$  is called negatively associated (NA) if for every pair of disjoint subsets  $S, T$  of  $Z$  and any real coordinatewise increasing functions  $f$  on  $R^S$  and  $g$  on  $R^T$

$$\text{Cov}\{f(Y_i, i \in S), g(Y_j, j \in T)\} \leq 0.$$

**Definition 1.4** A sequence  $\{Y_i, -\infty < i < \infty\}$  of random variables is said to be stochastically dominated by a random variable  $Y$  if there exists a constant  $C$  such that

$$P\{|Y_i| > x\} \leq CP\{|Y| > x\}, \quad x \geq 0, -\infty < i < \infty.$$

**Definition 1.5** A real valued function  $l(x)$ , positive and measurable on  $[0, \infty)$ , is said to be slowly varying at infinity if for each  $\lambda > 0$ ,  $\lim_{x \rightarrow \infty} \frac{l(\lambda x)}{l(x)} = 1$ .

Li and Zhang [11] obtained the following complete moment convergence of moving average processes under NA assumptions.

**Theorem A** Suppose that  $\{X_n = \sum_{i=-\infty}^{\infty} a_i \varepsilon_{i+n}, n \geq 1\}$ , where  $\{a_i, -\infty < i < \infty\}$  is a sequence of real numbers with  $\sum_{i=-\infty}^{\infty} |a_i| < \infty$  and  $\{\varepsilon_i, -\infty < i < \infty\}$  is a sequence of identically distributed NA random variables with  $E\varepsilon_1 = 0, E\varepsilon_1^2 < \infty$ . Let  $h$  be a function slowly varying at infinity,  $1 \leq q < 2, r > 1 + q/2$ . Then  $E|\varepsilon_1|^r h(|\varepsilon_1|^q) < \infty$  implies

$$\sum_{n=1}^{\infty} n^{r/q-2-1/q} h(n) E \left\{ \left| \sum_{j=1}^n X_j \right| - \varepsilon n^{1/q} \right\}^+ < \infty$$

for all  $\varepsilon > 0$ .

Chen *et al.* [20] also established the following results for moving average processes under NA assumptions.

**Theorem B** Let  $q > 0, 1 \leq p < 2, r \geq 1, rp \neq 1$ . Suppose that  $\{X_n = \sum_{i=-\infty}^{\infty} a_i \varepsilon_{i+n}, n \geq 1\}$ , where  $\{a_i, -\infty < i < \infty\}$  is a sequence of real numbers with  $\sum_{i=-\infty}^{\infty} |a_i| < \infty$  and  $\{\varepsilon_i, -\infty < i <$

$\infty$  is a sequence of identically distributed NA random variables. If  $E\varepsilon_1 = 0$  and  $E|\varepsilon_1|^p < \infty$ , then

$$\sum_{n=1}^{\infty} n^{r-2} P \left\{ \max_{1 \leq k \leq n} \left| \sum_{j=1}^k X_j \right| \geq \varepsilon n^{1/p} \right\} < \infty$$

for all  $\varepsilon > 0$ . Furthermore if  $E\varepsilon_1 = 0$  and  $E|\varepsilon_1|^p < \infty$  for  $q < rp$ ,  $E|\varepsilon_1|^p \log(1 + |\varepsilon_1|) < \infty$  for  $q = rp$ ,  $E|\varepsilon_1|^q < \infty$  for  $q > rp$ , then

$$\sum_{n=1}^{\infty} n^{r-2-q/p} E \left( \left\{ \max_{1 \leq k \leq n} \left| \sum_{j=1}^k X_j \right| - \varepsilon n^{1/q} \right\}^+ \right)^q < \infty$$

for all  $\varepsilon > 0$ .

Recently, Zhou and Lin [18] obtained the following complete moment convergence of moving average processes under  $\rho$ -mixing assumptions.

**Theorem C** *Let  $h$  be a function slowly varying at infinity,  $p \geq 1$ ,  $p\alpha > 1$  and  $\alpha > 1/2$ . Suppose that  $\{X_n, n \geq 1\}$  is a moving average process based on a sequence  $\{Y_i, -\infty < i < \infty\}$  of identically distributed  $\rho$ -mixing random variables. If  $EY_1 = 0$  and  $E|Y_1|^{p+\delta} h(|Y_1|^{1/\alpha}) < \infty$  for some  $\delta > 0$ , then for all  $\varepsilon > 0$ ,*

$$\sum_{n=1}^{\infty} n^{p\alpha-2-\alpha} h(n) E \left\{ \max_{1 \leq k \leq n} \left| \sum_{j=1}^k X_j \right| - \varepsilon n^\alpha \right\}^+ < \infty$$

and

$$\sum_{n=1}^{\infty} n^{p\alpha-2} h(n) E \left\{ \sup_{k \geq n} \left| k^{-\alpha} \sum_{j=1}^k X_j \right| - \varepsilon \right\}^+ < \infty.$$

Obviously,  $\rho^-$ -mixing random variables include NA and  $\rho^*$ -mixing random variables, which have a lot of applications, their limit properties have aroused wide interest recently, and a lot of results have been obtained; we refer to Wang and Lu [21] for a Rosenthal-type moment inequality and weak convergence, Budzaba *et al.* [22, 23] for complete convergence for moving average process based on a  $\rho^-$ -mixing sequence, Tan *et al.* [24] for the almost sure central limit theorem. But there are few results on the complete moment convergence of moving average process based on a  $\rho^-$ -mixing sequence. Therefore, in this paper, we establish some results on the complete moment convergence for maximum partial sums with less restrictions. Throughout the sequel,  $C$  represents a positive constant although its value may change from one appearance to the next,  $I\{A\}$  denotes the indicator function of the set  $A$ .

## 2 Preliminary lemmas

In this section, we list some lemmas which will be useful to prove our main results.

**Lemma 2.1** (Zhou [17]) *If  $l$  is slowly varying at infinity, then*

- (1)  $\sum_{n=1}^m n^s l(n) \leq Cm^{s+1} l(m)$  for  $s > -1$  and positive integer  $m$ ,
- (2)  $\sum_{n=m}^\infty n^s l(n) \leq Cm^{s+1} l(m)$  for  $s < -1$  and positive integer  $m$ .

**Lemma 2.2** (Wang and Lu [21]) *For a positive real number  $q \geq 2$ , if  $\{X_n, n \geq 1\}$  is a sequence of  $\rho^-$ -mixing random variables, with  $EX_i = 0, E|X_i|^q < \infty$  for every  $i \geq 1$ , then for all  $n \geq 1$ , there is a positive constant  $C = C(q, \rho^-(\cdot))$  such that*

$$E\left(\max_{1 \leq j \leq n} \left| \sum_{i=1}^j X_i \right|^q\right) \leq C \left\{ \sum_{i=1}^n E|X_i|^q + \left( \sum_{i=1}^n EX_i^2 \right)^{\frac{q}{2}} \right\}.$$

**Lemma 2.3** (Wang et al. [25]) *Let  $\{X_n, n \geq 1\}$  be a sequence of random variables which is stochastically dominated by a random variable  $X$ . Then for any  $a > 0$  and  $b > 0$ ,*

$$E|X_n|^a I\{|X_n| \leq b\} \leq C[E|X|^a I\{|X| \leq b\} + b^a P(|X| > b)],$$

$$E|X_n|^a I\{|X_n| > b\} \leq CE|X|^a I\{|X| > b\}.$$

### 3 Main results and proofs

**Theorem 3.1** *Let  $l$  be a function slowly varying at infinity,  $p \geq 1, \alpha > 1/2, \alpha p > 1$ . Assume that  $\{a_i, -\infty < i < \infty\}$  is an absolutely summable sequence of real numbers. Suppose that  $\{X_n = \sum_{i=-\infty}^\infty a_i Y_{i+n}, n \geq 1\}$  is a moving average process generated by a sequence  $\{Y_i, -\infty < i < \infty\}$  of  $\rho^-$ -mixing random variables which is stochastically dominated by a random variable  $Y$ . If  $EY_i = 0$  for  $1/2 < \alpha \leq 1, E|Y|^p l(|Y|^{1/\alpha}) < \infty$  for  $p > 1$  and  $E|Y|^{1+\delta} < \infty$  for  $p = 1$  and some  $\delta > 0$ , then for any  $\varepsilon > 0$*

$$\sum_{n=1}^\infty n^{\alpha p - 2 - \alpha} l(n) E \left\{ \max_{1 \leq k \leq n} \left| \sum_{j=1}^k X_j \right| - \varepsilon n^\alpha \right\}^+ < \infty \tag{3.1}$$

and

$$\sum_{n=1}^\infty n^{\alpha p - 2} l(n) E \left\{ \sup_{k \geq n} \left| k^{-\alpha} \sum_{j=1}^k X_j \right| - \varepsilon \right\}^+ < \infty. \tag{3.2}$$

*Proof* Firstly to prove (3.1). Let  $f(n) = n^{\alpha p - 2 - \alpha} l(n)$  and  $Y_{xj}^{(1)} = -xI\{Y_j < -x\} + Y_j I\{|Y_j| \leq x\} + xI\{Y_j > x\}$  and  $Y_{xj}^{(2)} = Y_j - Y_{xj}^{(1)}$  be the monotone truncations of  $\{Y_j, -\infty < j < \infty\}$  for  $x > 0$ . Then by the property of  $\rho^-$ -mixing random variables (cf. Property P2 in Wang and Lu [21]),  $\{Y_{xj}^{(1)} - EY_{xj}^{(1)}, -\infty < j < \infty\}$  and  $\{Y_{xj}^{(2)}, -\infty < j < \infty\}$  are two sequences of  $\rho^-$ -mixing random variables. Note that  $\sum_{k=1}^n X_k = \sum_{i=-\infty}^\infty a_i \sum_{j=i+1}^{i+n} Y_j$ . Since  $\sum_{i=-\infty}^\infty |a_i| < \infty$ , by Lemma 2.3, we have for  $x > n^\alpha$ , if  $\alpha > 1$

$$\begin{aligned} & x^{-1} \left| E \sum_{i=-\infty}^\infty a_i \sum_{j=i+1}^{i+n} Y_{xj}^{(1)} \right| \\ & \leq x^{-1} \sum_{i=-\infty}^\infty |a_i| \sum_{j=i+1}^{i+n} [E|Y_j| I\{|Y_j| \leq x\} + xP(|Y| > x)] \\ & \leq Cx^{-1} n [E|Y| I\{|Y| \leq x\} + xP(|Y| > x)] \leq Cn^{1-\alpha} \rightarrow 0, \quad \text{as } n \rightarrow \infty. \end{aligned}$$

If  $1/2 < \alpha \leq 1$ , note  $\alpha p > 1$ , this means  $p > 1$ . By  $E|Y|^p I(|Y|^{1/\alpha}) < \infty$  and  $l$  is slowly varying at infinity, for any  $0 < \epsilon < p - 1/\alpha$ , we have  $E|Y|^{p-\epsilon} < \infty$ . Then noting  $EY_i = 0$ , by Lemma 2.3 we have

$$\begin{aligned} x^{-1} \left| E \sum_{i=-\infty}^{\infty} a_i \sum_{j=i+1}^{i+n} Y_{xj}^{(1)} \right| &= x^{-1} \left| E \sum_{i=-\infty}^{\infty} a_i \sum_{j=i+1}^{i+n} Y_{xj}^{(2)} \right| \\ &\leq Cx^{-1} \sum_{i=-\infty}^{\infty} |a_i| \sum_{j=i+1}^{i+n} E|Y_j| I\{|Y_j| > x\} \leq Cx^{-1} n E|Y| I\{|Y| > x\} \\ &\leq Cx^{1/\alpha-1} E|Y| I\{|Y| > x\} \leq CE|Y|^{1/\alpha} I\{|Y| > x\} \\ &\leq E|Y|^{p-\epsilon} I\{|Y| > x\} \rightarrow 0, \quad \text{as } x \rightarrow \infty. \end{aligned}$$

Hence for  $x > n^\alpha$  large enough, we get

$$x^{-1} \left| E \sum_{i=-\infty}^{\infty} a_i \sum_{j=i+1}^{i+n} Y_{xj}^{(1)} \right| < \epsilon/4.$$

Therefore

$$\begin{aligned} &\sum_{n=1}^{\infty} f(n) E \left\{ \max_{1 \leq k \leq n} \left| \sum_{j=1}^k X_j \right| - \epsilon n^\alpha \right\}^+ \\ &\leq \sum_{n=1}^{\infty} f(n) \int_{\epsilon n^\alpha}^{\infty} P \left\{ \max_{1 \leq k \leq n} \left| \sum_{j=1}^k X_j \right| \geq x \right\} dx \\ &\leq C \sum_{n=1}^{\infty} f(n) \int_{n^\alpha}^{\infty} P \left\{ \max_{1 \leq k \leq n} \left| \sum_{j=1}^k X_j \right| \geq \epsilon x \right\} dx \\ &\leq C \sum_{n=1}^{\infty} f(n) \int_{n^\alpha}^{\infty} P \left\{ \max_{1 \leq k \leq n} \left| \sum_{i=-\infty}^{\infty} a_i \sum_{j=i+1}^{i+k} Y_{xj}^{(2)} \right| \geq \epsilon x/2 \right\} dx \\ &\quad + C \sum_{n=1}^{\infty} f(n) \int_{n^\alpha}^{\infty} P \left\{ \max_{1 \leq k \leq n} \left| \sum_{i=-\infty}^{\infty} a_i \sum_{j=i+1}^{i+k} (Y_{xj}^{(1)} - EY_{xj}^{(1)}) \right| \geq \epsilon x/4 \right\} dx \\ &=: I_1 + I_2. \tag{3.3} \end{aligned}$$

Firstly we show  $I_1 < \infty$ . Noting  $|Y_{xj}^{(2)}| < |Y_j| I\{|Y_j| > x\}$ , by Markov's inequality and Lemma 2.3, we have

$$\begin{aligned} I_1 &\leq C \sum_{n=1}^{\infty} f(n) \int_{n^\alpha}^{\infty} x^{-1} E \max_{1 \leq k \leq n} \left| \sum_{i=-\infty}^{\infty} a_i \sum_{j=i+1}^{i+k} Y_{xj}^{(2)} \right| dx \\ &\leq C \sum_{n=1}^{\infty} f(n) \int_{n^\alpha}^{\infty} x^{-1} \sum_{i=-\infty}^{\infty} |a_i| \sum_{j=i+1}^{i+n} E|Y_{xj}^{(2)}| dx \\ &\leq C \sum_{n=1}^{\infty} n f(n) \int_{n^\alpha}^{\infty} x^{-1} E|Y| I\{|Y| > x\} dx \end{aligned}$$

$$\begin{aligned}
 &= C \sum_{n=1}^{\infty} n f(n) \sum_{m=n}^{\infty} \int_{m^{\alpha}}^{(m+1)^{\alpha}} x^{-1} E|Y|I\{|Y| > x\} dx \\
 &\leq C \sum_{n=1}^{\infty} n f(n) \sum_{m=n}^{\infty} m^{-1} E|Y|I\{|Y| > m^{\alpha}\} \\
 &= C \sum_{m=1}^{\infty} m^{-1} E|Y|I\{|Y| > m^{\alpha}\} \sum_{n=1}^m n^{\alpha p-1-\alpha} l(n).
 \end{aligned}$$

If  $p > 1$ , then  $\alpha p - 1 - \alpha > -1$ , and, by Lemma 2.1, we obtain

$$\begin{aligned}
 I_1 &\leq C \sum_{m=1}^{\infty} m^{\alpha p-1-\alpha} l(m) E|Y|I\{|Y| > m^{\alpha}\} \\
 &= C \sum_{m=1}^{\infty} m^{\alpha p-1-\alpha} l(m) \sum_{k=m}^{\infty} E|Y|I\{k^{\alpha} < |Y| \leq (k+1)^{\alpha}\} \\
 &= C \sum_{k=1}^{\infty} E|Y|I\{k^{\alpha} < |Y| \leq (k+1)^{\alpha}\} \sum_{m=1}^k m^{\alpha p-1-\alpha} l(m) \\
 &\leq C \sum_{k=1}^{\infty} k^{\alpha p-\alpha} l(k) E|Y|I\{k^{\alpha} < |Y| \leq (k+1)^{\alpha}\} \\
 &\leq CE|Y|^p l(|Y|^{1/\alpha}) < \infty.
 \end{aligned}$$

If  $p = 1$ , notice that  $E|Y|^{1+\delta} < \infty$  implies  $E|Y|^{1+\delta'} l(|Y|^{1/\alpha}) < \infty$  for any  $0 < \delta' < \delta$ , then by Lemma 2.1, we obtain

$$\begin{aligned}
 I_1 &\leq C \sum_{m=1}^{\infty} m^{-1} E|Y|I\{|Y| > m^{\alpha}\} \sum_{n=1}^m n^{-1} l(n) \\
 &\leq C \sum_{m=1}^{\infty} m^{-1} E|Y|I\{|Y| > m^{\alpha}\} \sum_{n=1}^m n^{-1+\alpha\delta'} l(n) \\
 &\leq C \sum_{m=1}^{\infty} m^{\alpha\delta'-1} l(m) E|Y|I\{|Y| > m^{\alpha}\} \\
 &\leq CE|Y|^{1+\delta'} l(|Y|^{1/\alpha}) \leq CE|Y|^{1+\delta} < \infty.
 \end{aligned}$$

So, we get

$$I_1 < \infty. \tag{3.4}$$

Next we show  $I_2 < \infty$ . By Markov's inequality, the Hölder inequality, and Lemma 2.2, we conclude

$$\begin{aligned}
 I_2 &\leq C \sum_{n=1}^{\infty} f(n) \int_{n^{\alpha}}^{\infty} x^{-r} E \max_{1 \leq k \leq n} \left| \sum_{i=-\infty}^{\infty} a_i \sum_{j=i+1}^{i+k} (Y_{xj}^{(1)} - EY_{xj}^{(1)}) \right|^r dx \\
 &\leq C \sum_{n=1}^{\infty} f(n) \int_{n^{\alpha}}^{\infty} x^{-r} E \left[ \sum_{i=-\infty}^{\infty} (|a_i|^{\frac{r-1}{r}}) \left( |a_i|^{1/r} \max_{1 \leq k \leq n} \left| \sum_{j=i+1}^{i+k} (Y_{xj}^{(1)} - EY_{xj}^{(1)}) \right| \right) \right]^r dx
 \end{aligned}$$

$$\begin{aligned}
 &\leq C \sum_{n=1}^{\infty} f(n) \int_{n^\alpha}^{\infty} x^{-r} \left( \sum_{i=-\infty}^{\infty} |a_i| \right)^{r-1} \left( \sum_{i=-\infty}^{\infty} |a_i| E \max_{1 \leq k \leq n} \left| \sum_{j=i+1}^{i+k} (Y_{xj}^{(1)} - EY_{xj}^{(1)}) \right|^r \right) dx \\
 &\leq C \sum_{n=1}^{\infty} f(n) \int_{n^\alpha}^{\infty} x^{-r} \sum_{i=-\infty}^{\infty} |a_i| \sum_{j=i+1}^{i+n} E |Y_{xj}^{(1)} - EY_{xj}^{(1)}|^r dx \\
 &\quad + C \sum_{n=1}^{\infty} f(n) \int_{n^\alpha}^{\infty} x^{-r} \sum_{i=-\infty}^{\infty} |a_i| \left( \sum_{j=i+1}^{i+n} E |Y_{xj}^{(1)} - EY_{xj}^{(1)}|^2 \right)^{r/2} dx \\
 &=: I_{21} + I_{22},
 \end{aligned} \tag{3.5}$$

where  $r \geq 2$  will be specialized later.

For  $I_{21}$ , if  $p > 1$ , take  $r > \max\{2, p\}$ , then by  $C_r$  inequality, Lemma 2.3, and Lemma 2.1, we get

$$\begin{aligned}
 I_{21} &\leq C \sum_{n=1}^{\infty} f(n) \int_{n^\alpha}^{\infty} x^{-r} \sum_{i=-\infty}^{\infty} |a_i| \sum_{j=i+1}^{i+n} [E|Y_j|^r I\{|Y_j| \leq x\} + x^r P(|Y_j| > x)] dx \\
 &\leq C \sum_{n=1}^{\infty} n f(n) \int_{n^\alpha}^{\infty} x^{-r} [E|Y|^r I\{|Y| \leq x\} + x^r P(|Y| > x)] dx \\
 &\leq C \sum_{n=1}^{\infty} n f(n) \sum_{m=n}^{\infty} \int_{m^\alpha}^{(m+1)^\alpha} [x^{-r} E|Y|^r I\{|Y| \leq x\} + P(|Y| > x)] dx \\
 &\leq C \sum_{n=1}^{\infty} n f(n) \sum_{m=n}^{\infty} [m^{\alpha(1-r)-1} E|Y|^r I\{|Y| \leq (m+1)^\alpha\} + m^{\alpha-1} P(|Y| > m^\alpha)] \\
 &= C \sum_{m=1}^{\infty} [m^{\alpha(1-r)-1} E|Y|^r I\{|Y| \leq (m+1)^\alpha\} + m^{\alpha-1} P(|Y| > m^\alpha)] \sum_{n=1}^m n f(n) \\
 &\leq C \sum_{m=1}^{\infty} m^{\alpha(p-r)-1} l(m) \sum_{k=1}^m E|Y|^r I\{k^\alpha < |Y| \leq (k+1)^\alpha\} \\
 &\quad + C \sum_{m=1}^{\infty} m^{\alpha p-1} l(m) \sum_{k=m}^{\infty} E I\{k^\alpha < |Y| \leq (k+1)^\alpha\} \\
 &= C \sum_{k=1}^{\infty} E|Y|^r I\{k^\alpha < |Y| \leq (k+1)^\alpha\} \sum_{m=k}^{\infty} m^{\alpha(p-r)-1} l(m) \\
 &\quad + C \sum_{k=1}^{\infty} E I\{k^\alpha < |Y| \leq (k+1)^\alpha\} \sum_{m=1}^k m^{\alpha p-1} l(m) \\
 &\leq C \sum_{k=1}^{\infty} k^{\alpha(p-r)} l(k) E|Y|^p |Y|^{r-p} I\{k^\alpha < |Y| \leq (k+1)^\alpha\} \\
 &\quad + C \sum_{k=1}^{\infty} k^{\alpha p} l(k) E|Y|^p |Y|^{-p} I\{k^\alpha < |Y| \leq (k+1)^\alpha\} \\
 &\leq CE|Y|^p l(|Y|^{1/\alpha}) < \infty.
 \end{aligned} \tag{3.6}$$

For  $I_{21}$ , if  $p = 1$ , take  $r > \max\{1 + \delta', 2\}$ , where  $0 < \delta' < \delta$ , then by the same argument as above we have

$$\begin{aligned}
 I_{21} &\leq C \sum_{m=1}^{\infty} [m^{\alpha(1-r)-1} E|Y|^r I\{|Y| \leq (m+1)^\alpha\} + m^{\alpha-1} P(|Y| > m^\alpha)] \sum_{n=1}^m n f(n) \\
 &\leq C \sum_{m=1}^{\infty} [m^{\alpha(1-r)-1} E|Y|^r I\{|Y| \leq (m+1)^\alpha\} + m^{\alpha-1} P(|Y| > m^\alpha)] \sum_{n=1}^m n^{-1+\alpha\delta'} l(n) \\
 &\leq C \sum_{m=1}^{\infty} [m^{\alpha(1-r+\delta')-1} l(m) E|Y|^r I\{|Y| \leq (m+1)^\alpha\} \\
 &\quad + m^{\alpha(1+\delta')-1} l(m) E I\{|Y| > m^\alpha\}] \\
 &\leq C E|Y|^{1+\delta'} l(|Y|^{1/\alpha}) \leq C E|Y|^{1+\delta} < \infty.
 \end{aligned} \tag{3.7}$$

For  $I_{22}$ , if  $1 \leq p < 2$ , take  $r > 2$ , note  $\alpha p + r/2 - \alpha p r/2 - 1 = (\alpha p - 1)(1 - r/2) < 0$ , by the  $C_r$  inequality, Lemma 2.3, and Lemma 2.1, we obtain

$$\begin{aligned}
 I_{22} &\leq C \sum_{n=1}^{\infty} n^{r/2} f(n) \int_{n^\alpha}^{\infty} x^{-r} [(E|Y|^2 I\{|Y| \leq x\})^{r/2} + x^r P^{r/2}(|Y| > x)] dx \\
 &\leq C \sum_{n=1}^{\infty} n^{r/2} f(n) \sum_{m=n}^{\infty} \int_{m^\alpha}^{(m+1)^\alpha} [x^{-r} (E|Y|^2 I\{|Y| \leq x\})^{r/2} + P^{r/2}(|Y| > x)] dx \\
 &\leq C \sum_{n=1}^{\infty} n^{r/2} f(n) \sum_{m=n}^{\infty} [m^{\alpha(1-r)-1} (E|Y|^2 I\{|Y| \leq (m+1)^\alpha\})^{r/2} + m^{\alpha-1} P^{r/2}(|Y| > m^\alpha)] \\
 &= C \sum_{m=1}^{\infty} [m^{\alpha(1-r)-1} (E|Y|^2 I\{|Y| \leq (m+1)^\alpha\})^{r/2} + m^{\alpha-1} P^{r/2}(|Y| > m^\alpha)] \sum_{n=1}^m n^{r/2} f(n) \\
 &\leq C \sum_{m=1}^{\infty} m^{\alpha(p-r)+r/2-2} l(m) (E|Y|^p |Y|^{2-p} I\{|Y| \leq (m+1)^\alpha\})^{r/2} \\
 &\quad + C \sum_{m=1}^{\infty} m^{\alpha p+r/2-2} l(m) (E|Y|^p |Y|^{-p} I\{|Y| > m^\alpha\})^{r/2} \\
 &\leq C \sum_{m=1}^{\infty} m^{\alpha p+r/2-\alpha p r/2-2} l(m) (E|Y|^p)^{r/2} < \infty.
 \end{aligned} \tag{3.8}$$

For  $I_{22}$ , if  $p \geq 2$ , take  $r > (\alpha p - 1)/(\alpha - 1/2) > 2$ ; we have  $\alpha(p - r) + r/2 - 2 < -1$ , and therefore one gets

$$\begin{aligned}
 I_{22} &\leq C \sum_{m=1}^{\infty} [m^{\alpha(1-r)-1} (E|Y|^2 I\{|Y| \leq (m+1)^\alpha\})^{r/2} + m^{\alpha-1} P^{r/2}(|Y| > m^\alpha)] \sum_{n=1}^m n^{r/2} f(n) \\
 &\leq C \sum_{m=1}^{\infty} m^{\alpha(p-r)+r/2-2} l(m) (E|Y|^2 I\{|Y| \leq (m+1)^\alpha\})^{r/2} \\
 &\quad + C \sum_{m=1}^{\infty} m^{\alpha p+r/2-2} l(m) (E|Y|^2 |Y|^{-2} I\{|Y| > m^\alpha\})^{r/2}
 \end{aligned}$$



$$\leq C \sum_{m=1}^{\infty} m^{\alpha(p-r)+r/2-2} l(m) (E|Y|^2)^{r/2} < \infty. \tag{3.9}$$

Thus, (3.1) can be deduced by combining (3.3)-(3.9).

Now, we show (3.2). By Lemma 2.1 and (3.1) we have

$$\begin{aligned} & \sum_{n=1}^{\infty} n^{\alpha p-2} l(n) E \left\{ \sup_{k \geq n} \left| k^{-\alpha} \sum_{j=1}^k X_j \right| - \varepsilon \right\}^+ \\ &= \sum_{n=1}^{\infty} n^{\alpha p-2} l(n) \int_0^{\infty} P \left\{ \sup_{k \geq n} \left| k^{-\alpha} \sum_{j=1}^k X_j \right| > \varepsilon + t \right\} dt \\ &= \sum_{i=1}^{\infty} \sum_{n=2^{i-1}}^{2^i-1} n^{\alpha p-2} l(n) \int_0^{\infty} P \left\{ \sup_{k \geq n} \left| k^{-\alpha} \sum_{j=1}^k X_j \right| > \varepsilon + t \right\} dt \\ &\leq C \sum_{i=1}^{\infty} \int_0^{\infty} P \left\{ \sup_{k \geq 2^{i-1}} \left| k^{-\alpha} \sum_{j=1}^k X_j \right| > \varepsilon + t \right\} dt \sum_{n=2^{i-1}}^{2^i-1} n^{\alpha p-2} l(n) \\ &\leq C \sum_{i=1}^{\infty} 2^{i(\alpha p-1)} l(2^i) \int_0^{\infty} P \left\{ \sup_{k \geq 2^{i-1}} \left| k^{-\alpha} \sum_{j=1}^k X_j \right| > \varepsilon + t \right\} dt \\ &\leq C \sum_{i=1}^{\infty} 2^{i(\alpha p-1)} l(2^i) \sum_{l=i}^{\infty} \int_0^{\infty} P \left\{ \max_{2^{l-1} \leq k < 2^l} \left| k^{-\alpha} \sum_{j=1}^k X_j \right| > \varepsilon + t \right\} dt \\ &\leq C \sum_{l=1}^{\infty} \int_0^{\infty} P \left\{ \max_{2^{l-1} \leq k < 2^l} \left| k^{-\alpha} \sum_{j=1}^k X_j \right| > \varepsilon + t \right\} dt \sum_{i=1}^l 2^{i(\alpha p-1)} l(2^i) \\ &\leq C \sum_{l=1}^{\infty} 2^{l(\alpha p-1)} l(2^l) \int_0^{\infty} P \left\{ \max_{2^{l-1} \leq k < 2^l} \left| \sum_{j=1}^k X_j \right| > (\varepsilon + t) 2^{(l-1)\alpha} \right\} dt \\ &\leq C \sum_{l=1}^{\infty} 2^{l(\alpha p-1-\alpha)} l(2^l) \int_0^{\infty} P \left\{ \max_{1 \leq k < 2^l} \left| \sum_{j=1}^k X_j \right| > \varepsilon 2^{(l-1)\alpha} + y \right\} dy \\ &\leq C \sum_{n=1}^{\infty} n^{\alpha p-2-\alpha} l(n) \int_0^{\infty} P \left\{ \max_{1 \leq k < n} \left| \sum_{j=1}^k X_j \right| > \varepsilon n^{\alpha} 2^{-\alpha} + y \right\} dy \\ &= C \sum_{n=1}^{\infty} n^{\alpha p-2-\alpha} l(n) E \left\{ \max_{1 \leq k < n} \left| \sum_{j=1}^k X_j \right| - \varepsilon_0 n^{\alpha} \right\}^+ < \infty. \end{aligned}$$

Hence the proof of Theorem 3.1 is completed. □

The next theorem treats the case  $\alpha p = 1$ .

**Theorem 3.2** *Let  $l$  be a function slowly varying at infinity,  $1 \leq p < 2$ . Assume that  $\sum_{i=-\infty}^{\infty} |a_i|^{\theta} < \infty$ , where  $\theta$  belong to  $(0, 1)$  if  $p = 1$  and  $\theta = 1$  if  $1 < p < 2$ . Suppose that  $\{X_n = \sum_{i=-\infty}^{\infty} a_i Y_{i+n}, n \geq 1\}$  is a moving average process generated by a sequence  $\{Y_i, -\infty < i < \infty\}$  of  $\rho^-$ -mixing random variables which is stochastically dominated by a random variable  $Y$ .*

If  $EY_i = 0$  and  $E|Y|^p I(|Y|^p) < \infty$ , then for any  $\varepsilon > 0$

$$\sum_{n=1}^{\infty} n^{-1-1/p} l(n) E \left\{ \max_{1 \leq k \leq n} \left| \sum_{j=1}^k X_j \right| - \varepsilon n^{1/p} \right\}^+ < \infty. \tag{3.10}$$

*Proof* Let  $g(n) = n^{-1-1/p} l(n)$ . Similarly to the proof of (3.3), we have

$$\begin{aligned} & \sum_{n=1}^{\infty} g(n) E \left\{ \max_{1 \leq k \leq n} \left| \sum_{j=1}^k X_j \right| - \varepsilon n^{1/p} \right\}^+ \\ & \leq C \sum_{n=1}^{\infty} g(n) \int_{n^{1/p}}^{\infty} P \left\{ \max_{1 \leq k \leq n} \left| \sum_{i=-\infty}^{\infty} a_i \sum_{j=i+1}^{i+k} Y_{xj}^{(2)} \right| \geq \varepsilon x/2 \right\} dx \\ & \quad + C \sum_{n=1}^{\infty} g(n) \int_{n^{1/p}}^{\infty} P \left\{ \max_{1 \leq k \leq n} \left| \sum_{i=-\infty}^{\infty} a_i \sum_{j=i+1}^{i+k} (Y_{xj}^{(1)} - EY_{xj}^{(1)}) \right| \geq \varepsilon x/4 \right\} dx \\ & =: J_1 + J_2. \end{aligned} \tag{3.11}$$

For  $J_1$ , by Markov's inequality, the  $C_r$  inequality, Lemma 2.3, and Lemma 2.1, one gets

$$\begin{aligned} J_1 & \leq C \sum_{n=1}^{\infty} g(n) \int_{n^{1/p}}^{\infty} x^{-\theta} E \max_{1 \leq k \leq n} \left| \sum_{i=-\infty}^{\infty} a_i \sum_{j=i+1}^{i+k} Y_{xj}^{(2)} \right|^{\theta} dx \\ & \leq C \sum_{n=1}^{\infty} n g(n) \int_{n^{1/p}}^{\infty} x^{-\theta} E|Y|^{\theta} I\{|Y| > x\} dx \\ & = C \sum_{n=1}^{\infty} n g(n) \sum_{m=n}^{\infty} \int_{m^{1/p}}^{(m+1)^{1/p}} x^{-\theta} E|Y|^{\theta} I\{|Y| > x\} dx \\ & \leq C \sum_{n=1}^{\infty} n g(n) \sum_{m=n}^{\infty} m^{(1-\theta)/p-1} E|Y|^{\theta} I\{|Y| > m^{1/p}\} \\ & = C \sum_{m=1}^{\infty} m^{(1-\theta)/p-1} E|Y|^{\theta} I\{|Y| > m^{1/p}\} \sum_{n=1}^m n g(n) \\ & \leq C \sum_{m=1}^{\infty} m^{-\theta/p} l(m) E|Y|^{\theta} I\{|Y| > m^{1/p}\} \\ & = C \sum_{m=1}^{\infty} m^{-\theta/p} l(m) \sum_{k=m}^{\infty} E|Y|^{\theta} I\{k^{1/p} < |Y| < (k+1)^{1/p}\} \\ & = C \sum_{k=1}^{\infty} E|Y|^{\theta} I\{k^{1/p} < |Y| < (k+1)^{1/p}\} \sum_{m=1}^k m^{-\theta/p} l(m) \\ & \leq C \sum_{k=1}^{\infty} k^{1-\theta/p} l(k) E|Y|^{\theta} I\{k^{1/p} < |Y| < (k+1)^{1/p}\} \\ & \leq CE|Y|^{\theta} I(|Y|^p) < \infty. \end{aligned} \tag{3.12}$$

For  $J_2$ , similar to the proof of  $I_2$ , take  $r = 2$ , by Lemma 2.2, Lemma 2.3, and Lemma 2.1, we conclude

$$\begin{aligned}
 J_2 &\leq C \sum_{n=1}^{\infty} g(n) \int_{n^{1/p}}^{\infty} x^{-2} E \max_{1 \leq k \leq n} \left| \sum_{i=-\infty}^{\infty} a_i \sum_{j=i+1}^{i+k} (Y_{xj}^{(1)} - EY_{xj}^{(1)}) \right|^2 dx \\
 &\leq C \sum_{n=1}^{\infty} ng(n) \int_{n^{1/p}}^{\infty} x^{-2} [E|Y|^2 I\{|Y| \leq x\} + x^2 P(|Y| > x)] dx \\
 &= C \sum_{n=1}^{\infty} ng(n) \sum_{m=n}^{\infty} \int_{m^{1/p}}^{(m+1)^{1/p}} x^{-2} [E|Y|^2 I\{|Y| \leq x\} + x^2 P(|Y| > x)] dx \\
 &\leq C \sum_{n=1}^{\infty} ng(n) \sum_{m=n}^{\infty} [m^{-1-1/p} E|Y|^2 I\{|Y| \leq (m+1)^{1/p}\} + m^{1/p-1} P(|Y| > m^{1/p})] \\
 &= C \sum_{m=1}^{\infty} [m^{-1-1/p} E|Y|^2 I\{|Y| \leq (m+1)^{1/p}\} + m^{1/p-1} P(|Y| > m^{1/p})] \sum_{n=1}^m ng(n) \\
 &\leq C \sum_{m=1}^{\infty} [m^{-2/p} l(m) E|Y|^2 I\{|Y| \leq (m+1)^{1/p}\} + l(m) P(|Y| > m^{1/p})] \\
 &\leq CE|Y|^p l(|Y|^p) < \infty.
 \end{aligned} \tag{3.13}$$

Hence from (3.11)-(3.13), (3.10) holds. □

For the complete convergence and strong law of large numbers, we have the following corollary from the above theorems immediately.

**Corollary 3.3** *Under the assumptions of Theorem 3.1, for any  $\varepsilon > 0$  we have*

$$\sum_{n=1}^{\infty} n^{\alpha p-2} l(n) P \left\{ \max_{1 \leq k \leq n} \left| \sum_{j=1}^k X_j \right| > \varepsilon n^{\alpha} \right\} < \infty. \tag{3.14}$$

*Under the assumptions of Theorem 3.2, for any  $\varepsilon > 0$  we have*

$$\sum_{n=1}^{\infty} n^{-1} l(n) P \left\{ \max_{1 \leq k \leq n} \left| \sum_{j=1}^k X_j \right| > \varepsilon n^{1/p} \right\} < \infty; \tag{3.15}$$

*in particular, the assumptions  $EY_i = 0$  and  $E|Y|^p < \infty$  imply the following Marcinkiewicz-Zygmund strong law of large numbers:*

$$\lim_{n \rightarrow \infty} \frac{1}{n^{1/p}} \sum_{j=1}^n X_j = 0 \quad a.s. \tag{3.16}$$

**Remark 3.4** Corollary 3.3 provides complete convergence for the maximum of partial sums, which extends the corresponding results of Budsaba *et al.* [22, 23] and Theorem 1 of Baek *et al.* [1] with less restrictions. Since  $\rho^-$ -mixing random variables include NA and  $\rho^*$ -mixing random variables, our results also hold for NA and  $\rho^*$ -mixing, and therefore

Theorem 3.1 improves upon the above Theorem A from Li and Zhang [11] with less restrictions, and our results also extend and generalize the above Theorem B from Chen *et al.* [20] with  $q = 1$  partly.

**Remark 3.5** Obviously, the assumption that  $\{Y_i, -\infty < i < \infty\}$  is stochastically dominated by a random variable  $Y$  is weaker than the assumption of identical distribution of the random variables  $\{Y_i, -\infty < i < \infty\}$ , therefore the above results also hold for identically distributed random variables.

**Remark 3.6** Let  $a_0 = 1, a_i = 0, i \neq 0$ , then  $S_n = \sum_{k=1}^n X_k = \sum_{k=1}^n Y_k$ . Hence the above results hold when  $\{X_k, k \geq 1\}$  is a sequence of  $\rho^-$ -mixing random variables which is stochastically dominated by a random variable  $Y$ .

#### Competing interests

The author declares to have no competing interests.

#### Acknowledgements

This work was supported by National Natural Science Foundation of China (Grant No. 11101180) and the Science and Technology Development Program of Jilin Province (Grant No. 20130522096JH).

Received: 17 April 2015 Accepted: 23 July 2015 Published online: 08 August 2015

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