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On complete convergence for weighted sums of martingale-difference random fields

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Abstract

Let $\{a_{\mathbf{n},\mathbf{i}}, \mathbf{n} \in Z_{+}^d, \mathbf{i} \leq \mathbf{n}\}$ be an array of real numbers, and let $\{X_{\mathbf{i}}, \mathbf{i} \in Z_{+}^d\}$ be the martingale differences with respect to $\{\mathcal{G}_{\mathbf{n}}, \mathbf{n} \in Z_{+}^d\}$ satisfying $E(E(X|\mathcal{G}_k)|\mathcal{G}_{\mathbf{m}}) = E(X|\mathcal{G}_{k\wedge \mathbf{m}})$ a.s., where $\mathbf{k} \wedge \mathbf{m}$ denotes componentwise minimum, $\{\mathcal{G}_{\mathbf{k}}, \mathbf{k} \in Z_{+}^d\}$ is a family of σ -algebras such that $\forall \mathbf{k} \leq \mathbf{n}, \mathcal{G}_{\mathbf{k}} \subset \mathcal{G}_{\mathbf{n}} \subset \mathcal{G}$, and X is any integrable random variable defined on the initial probability space. The aim of this paper is to obtain some results concerning complete convergence of weighted sums $\sum_{\mathbf{i} \leq \mathbf{n}} a_{\mathbf{n},\mathbf{i}} X_{\mathbf{i}}$. **MSC:** 60F05; 60F15

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1 Introduction

The concept of complete convergence for sums of independent and identically distributed random variables was introduced by Hsu and Robbins [1] as follows: A sequence of random variables $\{X_n\}$ is said to be completely to a constant *c* if

$$\sum_{n=1}^{\infty} P(|X_n-c| > \epsilon) < \infty \quad \text{for all } \epsilon > 0.$$

This result has been generalized and extended to the random fields $\{X_n, n \in \mathbb{Z}_+^d\}$ of random variables. For example, Fazekas and Tómács [2] and Czerebak-Mrozowicz *et al.* [3] for fields of pairwise independent random variables, and Gut and Stadtmüller [4] for random fields of i.i.d. random variables.

Let Z_+ be the set of positive integers. For fixed $d \in Z_+$, set $Z_+^d = \{\mathbf{n} = (n_1, n_2, ..., n_d) : n_i \in Z_+, i = 1, 2, ..., d\}$ with coordinatewise partial order, \leq , *i.e.*, for $\mathbf{m} = (m_1, m_2, ..., m_d)$, $\mathbf{n} = (n_1, n_2, ..., n_d) \in Z_+^d$, $\mathbf{m} \leq \mathbf{n}$ if and only if $m_i \leq n_i$, i = 1, 2, ..., d. For $\mathbf{n} = (n_1, n_2, ..., n_d) \in Z_+^d$, let $|\mathbf{n}| = \prod_{i=1}^d n_i$. For a field $\{a_{\mathbf{n}}, \mathbf{n} \in Z_+^d\}$ of real numbers, the limit superior is defined by $\inf_{r \geq 1} \sup_{|\mathbf{n}| > r} a_{\mathbf{n}}$ and is denoted by $\limsup_{|\mathbf{n}| \to \infty} a_{\mathbf{n}}$.

Note that $|\mathbf{n}| \to \infty$ is equivalent to $\max\{n_1, n_2, \dots, n_d\} \to \infty$, which is weaker than the condition $\min\{n_1, n_2, \dots, n_d\} \to \infty$ when $d \ge 2$.

Let $\{X_{\mathbf{n}}, \mathbf{n} \in Z_{+}^{d}\}$ be a field of random variables, and let $\{a_{\mathbf{n},\mathbf{k}}, \mathbf{n} \in Z_{+}^{d}, \mathbf{k} \leq \mathbf{n}\}$ be an array of real numbers. The weighted sums $\sum_{\mathbf{k} \leq \mathbf{n}} a_{\mathbf{n},\mathbf{k}}X_{\mathbf{k}}$ can play an important role in various applied and theoretical problems, such as those of the least squares estimators (see Kafles

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and Bhaskara Rao [5]) and *M*-estimates (see Rao and Zhao [6]) in linear models, the nonparametric regression estimators (see Priestley and Chao [7]), *etc.* So, the study of the limiting behavior of the weighted sums is very important and significant (see Chen and Hao [8]).

Now, we consider the notion of martingale differences. Let $\{\mathcal{G}_{\mathbf{k}}, \mathbf{k} \in Z_{+}^{d}\}$ be a family of σ -algebras such that

$$\mathcal{G}_k \subset \mathcal{G}_n \subset \mathcal{G}, \quad \forall k \leq n,$$

and for any integrable random variable X defined on the initial probability space,

$$E(E(X|\mathcal{G}_{\mathbf{k}})|\mathcal{G}_{\mathbf{m}}) = E(X|\mathcal{G}_{\mathbf{k}\wedge\mathbf{m}}) \quad \text{a.s.,}$$

$$(1.1)$$

where $\mathbf{k} \wedge \mathbf{m}$ denotes the componentwise minimum.

An $\{\mathcal{G}_{\mathbf{k}}, \mathbf{k} \in Z^d_+\}$ -adapted, integrable process $\{Y_{\mathbf{k}}, \mathbf{k} \in Z^d_+\}$ is called a martingale if and only if

$$E(Y_{\mathbf{n}}|\mathcal{G}_{\mathbf{m}}) = Y_{\mathbf{m}\wedge\mathbf{n}}$$
 a.s.

Let us observe that for martingale $\{(Y_n, \mathcal{G}_n), n \in \mathbb{Z}_+^d\}$, the random variables

$$X_{\mathbf{n}} = \sum_{\mathbf{a} \in \{0,1\}^d} (-1)^{\sum_{i=1}^d a_i} Y_{\mathbf{n}-\mathbf{a}},$$

where $\mathbf{a} = (a_1, a_2, ..., a_d)$ and $\mathbf{n} \in Z_+^d$, are martingale differences with respect to $\{\mathcal{G}_{\mathbf{n}}, \mathbf{n} \in Z_+^d\}$ (see Kuczmaszewska and Lagodowski [9]).

For the results concerning complete convergence for martingale arrays obtained in the one-dimensional case, we refer to Lagodowski and Rychlik [10], Elton [11], Lesigne and Volny [12], Stoica [13] and Ghosal and Chandra [14]. Recently, complete convergence for martingale difference random fields was proved by Kuczmaszewska and Lagodowski [9].

The aim of this paper is to obtain some results concerning complete convergence of weighted sums $\sum_{i \le n} a_{n,i}X_i$, where $\{a_{n,i}, n \in Z_+^d, i \le n\}$ is an array of real numbers, and $\{X_i, i \in Z_+^d\}$ is the martingale differences with respect to $\{\mathcal{G}_n, n \in Z_+^d\}$ satisfying (1.1).

2 Results

The following moment maximal inequality provides us a useful tool to prove the main results of this section (see Kuczmaszewska and Lagodowski [9]).

Lemma 2.1 Let $\{(Y_n, \mathcal{G}_n), n \in \mathbb{Z}_+^d\}$ be a martingale, and let $\{(X_n, \mathcal{G}_n), n \in \mathbb{Z}_+^d\}$ be the martingale differences corresponding to it. Let q > 1. There exists a finite and positive constant *C* depending only on *q* and *d* such that

$$E\left(\max_{\mathbf{k}\leq\mathbf{n}}|Y_{\mathbf{k}}|^{q}\right)\leq CE\left(\sum_{\mathbf{k}\leq\mathbf{n}}X_{\mathbf{k}}^{2}\right)^{q/2}.$$
(2.1)

Let us denote $\mathcal{G}_{i}^{*} = \sigma \{\mathcal{G}_{j} : j < i\}$. Now, we are ready to formulate the next result.

Theorem 2.2 Let $\{a_{\mathbf{n}}, \mathbf{i}, \mathbf{n} \in Z_{+}^{d}, \mathbf{i} \leq \mathbf{n}\}$ be an array of real numbers, and let $\{X_{\mathbf{n}}, \mathbf{n} \in Z_{+}^{d}\}$ be the martingale differences with respect to $\{\mathcal{G}_{\mathbf{n}}, \mathbf{n} \in Z_{+}^{d}\}$ satisfying (1.1). For $\alpha p > 1, p > 1$ and $\alpha > \frac{1}{2}$, we assume that

$$\begin{array}{ll} (i) & \sum_{\mathbf{n}} |\mathbf{n}|^{\alpha p-2} \sum_{\mathbf{i} \leq \mathbf{n}} P\{|a_{\mathbf{n},\mathbf{i}}X_{\mathbf{i}}| > |\mathbf{n}|^{\alpha}\} < \infty, \\ (ii) & \sum_{\mathbf{n}} |\mathbf{n}|^{\alpha (p-q)-3+q/2} \sum_{\mathbf{i} \leq \mathbf{n}} |a_{\mathbf{n},\mathbf{i}}|^{q} E(|X_{\mathbf{i}}|^{q} I[|a_{\mathbf{n},\mathbf{i}}X_{\mathbf{i}}| \leq |\mathbf{n}|^{\alpha}]) < \infty \ for \ q \geq 2, \\ (ii)' & \sum_{\mathbf{n}} |\mathbf{n}|^{\alpha (p-q)-2} \sum_{\mathbf{i} \leq \mathbf{n}} |a_{\mathbf{n},\mathbf{i}}|^{q} E(|X_{\mathbf{i}}|^{q} I[|a_{\mathbf{n},\mathbf{i}}X_{\mathbf{i}}| \leq |\mathbf{n}|^{\alpha}]) < \infty \ for \ 1 < q < 2 \ and \\ (iii) & \sum_{\mathbf{n}} |\mathbf{n}|^{\alpha p-2} P\{\max_{\mathbf{j} \leq \mathbf{n}} |\sum_{\mathbf{i} \leq \mathbf{j}} E(a_{\mathbf{n},\mathbf{i}}X_{\mathbf{i}}I[|a_{\mathbf{n},\mathbf{i}}X_{\mathbf{i}}| \leq |\mathbf{n}|^{\alpha}] |\mathcal{G}_{\mathbf{i}}^{*})| > \epsilon |\mathbf{n}|^{\alpha}\} < \infty \ for \ all \ \epsilon > 0. \end{array}$$

Then we have

$$\sum_{\mathbf{n}} |\mathbf{n}|^{\alpha p-2} P\left\{ \max_{\mathbf{j} \le \mathbf{n}} |S_{\mathbf{j}}| > \epsilon |\mathbf{n}|^{\alpha} \right\} < \infty \quad for \ all \ \epsilon > 0,$$
(2.2)

where $S_n = \sum_{1 \le i \le n} a_{n,i} X_i$.

Proof Let us notice that the series $\sum_{\mathbf{n}} |\mathbf{n}|^{\alpha p-2}$ is finite, then (2.2) always holds. Therefore, we consider only the case such that $\sum_{\mathbf{n}} |\mathbf{n}|^{\alpha p-2}$ is divergent. Let $X_{\mathbf{n},\mathbf{i}} = X_{\mathbf{i}}I[|a_{\mathbf{n},\mathbf{i}}X_{\mathbf{i}}| \le |\mathbf{n}|^{\alpha}]$, $X_{\mathbf{n},\mathbf{i}}^* = X_{\mathbf{n},\mathbf{i}} - E(X_{\mathbf{n},\mathbf{i}}|\mathcal{G}_{\mathbf{i}}^*)$ and $S_{\mathbf{n},\mathbf{j}}^* = \sum_{\mathbf{i}\le \mathbf{j}} a_{\mathbf{n},\mathbf{i}}X_{\mathbf{n},\mathbf{i}}^*$.

Then

$$\begin{split} \sum_{\mathbf{n}} |\mathbf{n}|^{\alpha p-2} P\Big\{ \max_{\mathbf{j} \leq \mathbf{n}} |S_{\mathbf{j}}| > \epsilon |\mathbf{n}|^{\alpha} \Big\} \\ &\leq \sum_{\mathbf{n}} |\mathbf{n}|^{\alpha p-2} P\Big\{ |a_{\mathbf{n},\mathbf{i}}X_{\mathbf{i}}| > |\mathbf{n}|^{\alpha} \Big\} \\ &+ \sum_{\mathbf{n}} |\mathbf{n}|^{\alpha p-2} P\Big\{ \max_{\mathbf{j} \leq \mathbf{n}} \Big| \sum_{\mathbf{i} \leq \mathbf{j}} a_{\mathbf{n},\mathbf{i}} X_{\mathbf{i}} I\Big[|a_{\mathbf{n},\mathbf{i}}X_{\mathbf{i}}| \le |\mathbf{n}|^{\alpha} \Big] \Big| > \epsilon |\mathbf{n}|^{\alpha} \Big\} \\ &\leq \sum_{\mathbf{n}} |\mathbf{n}|^{\alpha p-2} \sum_{\mathbf{i} \leq \mathbf{n}} P\Big\{ |a_{\mathbf{n},\mathbf{i}}X_{\mathbf{i}}| > |\mathbf{n}|^{\alpha} \Big\} + \sum_{\mathbf{n}} |\mathbf{n}|^{\alpha p-2} P\Big\{ \max_{\mathbf{j} \leq \mathbf{n}} \Big| \sum_{\mathbf{i} \leq \mathbf{j}} (a_{\mathbf{n},\mathbf{i}}X_{\mathbf{i}} I\Big[|a_{\mathbf{n},\mathbf{i}}X_{\mathbf{i}}| \le |\mathbf{n}|^{\alpha} \Big] \\ &- E\Big(a_{\mathbf{n},\mathbf{i}}X_{\mathbf{i}} I\Big[|a_{\mathbf{n},\mathbf{i}}X_{\mathbf{i}}| \le |\mathbf{n}|^{\alpha} \Big] |\mathcal{G}_{\mathbf{i}}^{*} \Big) \Big| > \frac{\epsilon}{2} |\mathbf{n}|^{\alpha} \Big\} \\ &+ \sum_{\mathbf{n}} |\mathbf{n}|^{\alpha p-2} P\Big\{ \max_{\mathbf{j} \leq \mathbf{n}} \Big| \sum_{\mathbf{i} \leq \mathbf{j}} E\Big(a_{\mathbf{n},\mathbf{i}}X_{\mathbf{i}} I\Big[|a_{\mathbf{n},\mathbf{i}}X_{\mathbf{i}}| \le |\mathbf{n}|^{\alpha} \Big] |\mathcal{G}_{\mathbf{i}}^{*} \Big) \Big| > \frac{\epsilon}{2} |\mathbf{n}|^{\alpha} \Big\} \\ &= I_{1} + I_{2} + I_{3}. \end{split}$$

Clearly, $I_1 < \infty$ by (i), and $I_3 < \infty$ by (iii). It remains to prove that $I_2 < \infty$. Thus, the proof will be completed by proving that

$$\sum_{\mathbf{n}} |\mathbf{n}|^{\alpha p-2} P\left\{ \max_{j \le \mathbf{n}} \left| S_{\mathbf{n},j}^* \right| > \epsilon |\mathbf{n}|^{\alpha} \right\} < \infty.$$

To prove it, we first observe that $\{(S_{n,j}^*, \mathcal{G}_j), j \leq n\}$ is a martingale. In fact, if i > j, then $\mathcal{G}_{i \wedge j} \subset \mathcal{G}_i^*$ and by (1.1), we have

$$E(a_{\mathbf{n},\mathbf{i}}X_{\mathbf{n},\mathbf{i}}^*|\mathcal{G}_{\mathbf{j}}) = E(a_{\mathbf{n},\mathbf{i}}X_{\mathbf{n},\mathbf{i}} - E(a_{\mathbf{n},\mathbf{i}}X_{\mathbf{n},\mathbf{i}}|\mathcal{G}_{\mathbf{i}}^*)|\mathcal{G}_{\mathbf{i}})$$
$$= E(E(a_{\mathbf{n},\mathbf{i}}X_{\mathbf{n},\mathbf{i}} - E(a_{\mathbf{n},\mathbf{i}}X_{\mathbf{n},\mathbf{i}}|\mathcal{G}_{\mathbf{i}}^*)|\mathcal{G}_{\mathbf{i}})|\mathcal{G}_{\mathbf{j}})$$

$$= E(a_{\mathbf{n},\mathbf{i}}X_{\mathbf{n},\mathbf{i}} - E(a_{\mathbf{n},\mathbf{i}}X_{\mathbf{n},\mathbf{i}}|\mathcal{G}_{\mathbf{i}}^*)|\mathcal{G}_{\mathbf{i}\wedge\mathbf{j}})$$
$$= 0.$$

Then, by the Markov inequality and Lemma 2.1, there exists some constant C such that

$$\begin{split} P\Big\{\max_{\mathbf{j}\leq\mathbf{n}} \left|S_{\mathbf{n},\mathbf{j}}^*\right| &> \epsilon \,|\mathbf{n}|^{\alpha}\Big\} \leq C \frac{E(\max_{\mathbf{j}\leq\mathbf{n}} |S_{\mathbf{n},\mathbf{j}}^*|^q)}{|\mathbf{n}|^{\alpha q}} \\ &\leq \frac{C}{|\mathbf{n}|^{\alpha q}} E\left(\sum_{\mathbf{i}\leq\mathbf{n}} a_{\mathbf{n},\mathbf{i}}^2 X_{\mathbf{n},\mathbf{i}}^{*2}\right)^{q/2} = I_4. \end{split}$$

Case $q \ge 2$; we get

$$I_4 \leq \frac{C}{|\mathbf{n}|^{\alpha q}} |\mathbf{n}|^{q/2-1} \sum_{\mathbf{i} \leq \mathbf{n}} E |a_{\mathbf{n},\mathbf{i}} X_{\mathbf{n},\mathbf{i}}^*|^q$$
$$\leq C |\mathbf{n}|^{q/2-1-\alpha q} \sum_{\mathbf{i} \leq \mathbf{n}} E (|a_{\mathbf{n},\mathbf{i}} X_{\mathbf{i}}|^q I [|a_{\mathbf{n},\mathbf{i}} X_{\mathbf{i}}| \leq |\mathbf{n}|^{\alpha}]).$$

Note that the last estimation follows from the Jensen inequality. Thus, we have

$$\sum_{\mathbf{n}} |\mathbf{n}|^{\alpha p-2} P\left\{ \max_{\mathbf{j} \le \mathbf{n}} \left| S_{\mathbf{n},\mathbf{j}}^* \right| > \epsilon |\mathbf{n}|^{\alpha} \right\}$$
$$\leq C \sum_{\mathbf{n}} |\mathbf{n}|^{\alpha p-3-q(\alpha-1/2)} \sum_{\mathbf{i} \le \mathbf{n}} E\left(|a_{\mathbf{n},\mathbf{i}}X_{\mathbf{i}}|^q I\left[|a_{\mathbf{n},\mathbf{i}}X_{\mathbf{i}}| \le |\mathbf{n}|^{\alpha} \right] \right) < \infty$$

by assumption (ii).

Case 1 < q < 2; we get

$$I_{4} \leq \frac{C}{|\mathbf{n}|^{\alpha q}} \sum_{\mathbf{i} \leq \mathbf{n}} E |a_{\mathbf{n},\mathbf{i}}X_{\mathbf{n},\mathbf{i}}^{*}|^{q}$$
$$\leq C|\mathbf{n}|^{-\alpha q} \sum_{\mathbf{i} \leq \mathbf{n}} E (|a_{\mathbf{n},\mathbf{i}}X_{\mathbf{i}}|^{q} I [|a_{\mathbf{n},\mathbf{i}}X_{\mathbf{i}}| \leq |\mathbf{n}|^{\alpha}]).$$

Therefore, for 1 < q < 2, we obtain

$$\sum_{\mathbf{n}} |\mathbf{n}|^{\alpha p-2} P\left\{ \max_{\mathbf{j} \le \mathbf{n}} \left| S_{\mathbf{n},\mathbf{j}}^* \right| > \epsilon |\mathbf{n}|^{\alpha} \right\}$$
$$\leq C \sum_{\mathbf{n}} |\mathbf{n}|^{\alpha (p-q)-2} \sum_{\mathbf{i} \le \mathbf{n}} E\left(|a_{\mathbf{n},\mathbf{i}} X_{\mathbf{i}}|^q I\left[|a_{\mathbf{n},\mathbf{i}} X_{\mathbf{i}}| \le |\mathbf{n}|^{\alpha} \right] \right) < \infty$$

by assumption (ii)'. Thus, $I_2 < \infty$ for all q > 1, and the proof of Theorem 2.2 is complete.

Corollary 2.3 Let $\{a_{\mathbf{n}}, \mathbf{i}, \mathbf{n} \in Z_{+}^{d}, \mathbf{i} \leq \mathbf{n}\}$ be an array of real numbers. Let $\{X_{\mathbf{n}}, \mathbf{n} \in Z_{+}^{d}\}$ be martingale differences with respect to $\{\mathcal{G}_{\mathbf{n}}, \mathbf{n} \in Z_{+}^{d}\}$ satisfying (1.1), and $EX_{\mathbf{n}} = 0$ for $\mathbf{n} \in Z_{+}^{d}$. Let $p \geq 1, \alpha > \frac{1}{2}$ and $\alpha p > 1$. Assume that (i) and for some q > 1, (ii) or (ii)' hold respectively. If

$$\max_{\mathbf{j}\leq\mathbf{n}}\sum_{\mathbf{i}\leq\mathbf{j}}E(a_{\mathbf{n},\mathbf{i}}X_{\mathbf{i}}I[|a_{\mathbf{n},\mathbf{i}}X_{\mathbf{i}}|\leq|\mathbf{n}|^{\alpha}]|\mathcal{G}_{\mathbf{i}}^{*})=o(|\mathbf{n}|^{\alpha}),$$
(2.3)

then (2.2) holds.

Proof It is easy to see that (2.3) implies (iii). We omit details that prove it.

The following corollary shows that assumption (iii) in Theorem 2.2 is natural, and in the case of independent random fields, it reduces to the known one.

Corollary 2.4 Let $\{a_{\mathbf{n}}, \mathbf{i}, \mathbf{n} \in \mathbb{Z}_{+}^{d}, \mathbf{i} \leq \mathbf{n}\}$ be an array of real numbers. Let $\{X_{\mathbf{n}}, \mathbf{n} \in \mathbb{Z}_{+}^{d}\}$ be a field of independent random variables such that $EX_{\mathbf{n}} = 0$ for $\mathbf{n} \in \mathbb{Z}_{+}^{d}$. Let $p \geq 1$, $\alpha > \frac{1}{2}$ and $\alpha p > 1$. Assume that (i) and for some q > 1, (ii) or (ii)' hold respectively. If

$$\frac{1}{|\mathbf{n}|^{\alpha}} \max_{j \le \mathbf{n}} \sum_{i \le j} E\left(a_{\mathbf{n},i} X_i I\left[\left|a_{\mathbf{n},i} X_i\right| \le \left|\mathbf{n}\right|^{\alpha}\right]\right) \to 0 \quad as \ |\mathbf{n}| \to \infty,$$
(2.4)

then (2.2) holds.

Proof Since $\{X_n, n \in \mathbb{Z}_+^d\}$ is a field of independent random variables, we have

$$\frac{1}{|\mathbf{n}|^{\alpha}} \max_{j \le \mathbf{n}} \sum_{i \le j} E\left(a_{\mathbf{n},i} X_i I\left[|a_{\mathbf{n},i} X_i| \le |\mathbf{n}|^{\alpha}\right]|\mathcal{G}_i^*\right) = \frac{1}{|\mathbf{n}|^{\alpha}} \max_{j \le \mathbf{n}} \sum_{i \le j} E\left(a_{\mathbf{n},i} X_i I\left[|a_{\mathbf{n},i} X_i| \le |\mathbf{n}|^{\alpha}\right]\right).$$

Now, it is easy to see that (2.4) implies (iii) of Theorem 2.2. Thus, by Theorem 2.2, result (2.2) follows. \Box

Remark Theorem 2.2 and Corollary 2.4 are extensions of Theorem 4.1 and Corollary 4.1 in Kuczmaszewska and Lagodowski [9] to the weighted sums case, respectively.

Corollary 2.5 Let $\{a_{\mathbf{n}}, \mathbf{i}, \mathbf{n} \in \mathbb{Z}_{+}^{d}, \mathbf{1} \leq \mathbf{i} \leq \mathbf{n}\}\$ be an array of real numbers. Let $\{X_{\mathbf{n}}, \mathbf{n} \in \mathbb{Z}_{+}^{d}\}\$ be the martingale differences with respect to $\{\mathcal{G}_{\mathbf{n}}, \mathbf{n} \in \mathbb{Z}_{+}^{d}\}\$ satisfying (1.1) and $EX_{\mathbf{n}} = 0$. Let $p \geq 1, \alpha > \frac{1}{2}$ and $\alpha p > 1$ and $E|X_{\mathbf{n}}|^{1+\lambda_{\mathbf{n}}} < \infty$ for $\lambda_{\mathbf{n}}$ with $0 < \lambda_{\mathbf{n}} < 1$ for $\mathbf{n} \in \mathbb{Z}_{+}^{d}$. If

$$\sum_{\mathbf{n}} |\mathbf{n}|^{\alpha p-2} |\mathbf{n}|^{-\alpha(1+\lambda_{\mathbf{n}})} \sum_{\mathbf{i} \le \mathbf{n}} |a_{\mathbf{n},\mathbf{i}}|^{1+\lambda_{\mathbf{n}}} E|X_{\mathbf{i}}|^{1+\lambda_{\mathbf{n}}} < \infty,$$
(2.5)

$$\max_{1 \le j \le \mathbf{n}} \sum_{\mathbf{i} \le j} E(|a_{\mathbf{n},\mathbf{i}}X_i| I[|a_{\mathbf{n},\mathbf{i}}X_{\mathbf{i}}| \le |\mathbf{n}|^{\alpha}]|\mathcal{G}_{\mathbf{i}}^*) = o(|\mathbf{n}|^{\alpha}),$$
(2.6)

then (2.2) holds.

Proof If the series $\sum_{\mathbf{n}} |\mathbf{n}|^{\alpha p-2} < \infty$, then (2.2) always holds. Hence, we only consider the case $\sum_{\mathbf{n}} |\mathbf{n}|^{\alpha p-2} = \infty$. It follows from (2.5) that

$$|\mathbf{n}|^{-\alpha(1+\lambda_{\mathbf{n}})}\sum_{\mathbf{i}\leq\mathbf{n}}|a_{\mathbf{n},\mathbf{i}}|^{1+\lambda_{\mathbf{n}}}E|X_{i}|^{1+\lambda_{\mathbf{n}}}<1.$$

By (2.5) and the Markov inequality,

$$\sum_{\mathbf{n}} |\mathbf{n}|^{\alpha p-2} P(|a_{\mathbf{n},\mathbf{i}}X_{\mathbf{i}}| > |\mathbf{n}|^{\alpha})$$

$$\leq \sum_{\mathbf{n}} |\mathbf{n}|^{\alpha p-2} |\mathbf{n}|^{-\alpha(1+\lambda_{\mathbf{n}})} \sum_{\mathbf{i} \le \mathbf{n}} |a_{\mathbf{n},\mathbf{i}}|^{1+\lambda_{\mathbf{n}}} E|X_{\mathbf{i}}|^{1+\lambda_{\mathbf{n}}} < \infty, \qquad (2.7)$$

which satisfies (i) of Theorem 2.2.

As the proof of Corollary 2.3, (2.6) implies (iii) of Theorem 2.2. It remains to show that Theorem 2.2(ii) or (ii)' is satisfied. For 1 < q < 2, take $1 + \lambda_n < q$. Then we have

$$\begin{split} &\sum_{\mathbf{n}} |\mathbf{n}|^{\alpha(p-q)-2} \sum_{\mathbf{i} \leq \mathbf{n}} |a_{\mathbf{n},\mathbf{i}}|^{q} E(|X_{\mathbf{i}}|^{q} I[|a_{\mathbf{n},\mathbf{i}}X_{\mathbf{i}}| \leq |\mathbf{n}|^{\alpha}]) \\ &\leq \sum_{\mathbf{n}} |\mathbf{n}|^{\alpha p-2} |\mathbf{n}|^{-\alpha(1+\lambda_{\mathbf{n}})} |\mathbf{n}|^{-\alpha q+\alpha(1+\lambda_{\mathbf{n}})} |\mathbf{n}|^{\alpha q-\alpha(1+\lambda_{\mathbf{n}})} \sum_{\mathbf{i} \leq \mathbf{n}} |a_{\mathbf{n},\mathbf{i}}|^{1+\lambda_{\mathbf{n}}} E|X_{\mathbf{i}}|^{1+\lambda_{\mathbf{n}}} \\ &= \sum_{\mathbf{n}} |\mathbf{n}|^{\alpha p-2} |\mathbf{n}|^{-\alpha(1+\lambda_{\mathbf{n}})} \sum_{\mathbf{i} \leq \mathbf{n}} |a_{\mathbf{n},\mathbf{i}}|^{1+\lambda_{\mathbf{n}}} E|X_{\mathbf{i}}|^{1+\lambda_{\mathbf{n}}} < \infty \quad \text{by (2.5),} \end{split}$$

which satisfies Theorem 2.2(ii)'. Hence, the proof is complete.

Corollary 2.6 Let $\{a_{\mathbf{n}}, \mathbf{i}, \mathbf{n} \in Z_{+}^{d}, \mathbf{1} \leq \mathbf{i} \leq \mathbf{n}\}$ be an array of real numbers, and let $\{X_{\mathbf{n}}, \mathbf{n} \in Z_{+}^{d}\}$ be the martingale differences with respect to $\{\mathcal{G}_{\mathbf{n}}, \mathbf{n} \in Z_{+}^{d}\}$ satisfying (1.1), $EX_{\mathbf{n}} = 0$ and $E|X_{\mathbf{n}}|^{p} < \infty$ for $1 . Let <math>\alpha > \frac{1}{2}$, $\alpha p > 1$ and 1 . If

$$\sum_{1 \le \mathbf{i} \le \mathbf{n}} |a_{\mathbf{n},\mathbf{i}}|^p E|X_{\mathbf{i}}|^p = O(|\mathbf{n}|^{\delta}) \quad for \ 0 < \delta < 1,$$
(2.8)

and Theorem 2.2(iii) hold, then (2.2) holds.

Proof By (2.8) and the Markov inequality,

$$\sum_{\mathbf{n}} |\mathbf{n}|^{\alpha p-2} \sum_{\mathbf{i} \le \mathbf{n}} P(|a_{\mathbf{n},\mathbf{i}}X_{\mathbf{i}}| > |\mathbf{n}|^{\alpha})$$

$$\leq \sum_{\mathbf{n}} |\mathbf{n}|^{\alpha p-2} \sum_{\mathbf{i} \le \mathbf{n}} \frac{|a_{\mathbf{n},\mathbf{i}}|^{p} E|X_{\mathbf{i}}|^{p}}{|\mathbf{n}|^{\alpha p}}$$

$$\leq C \sum_{\mathbf{n}} |\mathbf{n}|^{-2+\delta} < \infty.$$
(2.9)

By taking q < p, we have

$$\sum_{\mathbf{n}} |\mathbf{n}|^{\alpha(p-q)-2} \sum_{\mathbf{i} \le \mathbf{n}} |a_{\mathbf{n},\mathbf{i}}|^q E(|X_{\mathbf{i}}|^q I[|a_{\mathbf{n},\mathbf{i}}X_{\mathbf{i}}| \le \epsilon |\mathbf{n}|^{\alpha}])$$

$$\leq \sum_{\mathbf{n}} |\mathbf{n}|^{-2} \sum_{\mathbf{i} \le \mathbf{n}} |a_{\mathbf{n},\mathbf{i}}|^p E|X_{\mathbf{i}}|^p$$

$$\leq C \sum_{\mathbf{n}} |\mathbf{n}|^{-2+\delta} < \infty.$$
(2.10)

Hence, by (2.9) and (2.10), conditions (i) and (ii)' in Theorem 2.2 are satisfied, respectively.

To complete the proof, it is enough to note that by $EX_n = 0$ for $n \in Z^d_+$ and by (2.8), we get for $j \le n$

$$|\mathbf{n}|^{-\alpha} \sum_{\mathbf{i} \le \mathbf{j}} |a_{\mathbf{n},\mathbf{i}}| E|X_{\mathbf{i}}| I[|a_{\mathbf{n},\mathbf{i}}X_{\mathbf{i}}| \le \epsilon |\mathbf{n}|^{\alpha}] \to 0 \quad \text{as } |\mathbf{n}| \to \infty.$$

$$(2.11)$$

Hence, the proof is complete.

Corollary 2.7 Let $\{X_{\mathbf{n}}, \mathbf{n} \in Z_{+}^{d}\}$ be the martingale differences with respect to $\{\mathcal{G}_{\mathbf{n}}, \mathbf{n} \in Z_{+}^{d}\}$ satisfying (1.1), let $EX_{\mathbf{n}} = 0$ and $E|X_{\mathbf{n}}|^{p} < \infty$ for $1 and be stochastically dominated by a random variable X, i.e., there is a constant D such that <math>P(|X_{\mathbf{n}}| > x) \leq DP(|X| > x)$ for all $x \geq 0$ and $\mathbf{n} \in Z_{+}^{d}$. Let $\{a_{\mathbf{n}}, \mathbf{i}, \mathbf{n} \in Z_{+}^{d}, \mathbf{i} \leq \mathbf{n}\}$ be an array of real numbers satisfying

$$\sum_{\mathbf{i} \leq \mathbf{n}} |a_{\mathbf{n},\mathbf{i}}|^p = O(|\mathbf{n}|^{\delta}) \quad \text{for } 0 < \delta < 1.$$
(2.12)

If Theorem 2.2(iii) holds, then (2.2) holds.

Proof From (2.12), (2.8) follows. Hence, by Corollary 2.6, we obtain (2.2). \Box

Remark Linear random fields are of great importance in time series analysis. They arise in a wide variety of context. Applications to economics, engineering, and physical science are extremely broad (see Kim *et al.* [15]).

Let $Y_{\mathbf{k}} = \sum_{i \ge 1} a_{\mathbf{k}+i} X_i$, where $\{a_i, i \in \mathbb{Z}_+^d\}$ is a field of real numbers with $\sum_i \ge 1 |a_i| < \infty$, and $\{X_i, i \in \mathbb{Z}_+^d\}$ is a field of the martingale difference random variables.

Define $a_{\mathbf{n},\mathbf{i}} = \sum_{1 \le \mathbf{k} \le \mathbf{n}} a_{\mathbf{i}+\mathbf{k}}$. Then we have

$$\sum_{1 \le k \le n} Y_k = \sum_{1 \le k \le n} \sum_{i \ge 1} a_{i+k} X_i = \sum_{i \ge 1} \sum_{1 \le k \le n} a_{i+k} X_i = \sum_{i \ge 1} a_{n,i} X_i.$$

Hence, we can use the above results to investigate the complete convergence for linear random fields.

Competing interests

The author declares that she has no competing interests.

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